

Kernelized Cumulants: Beyond Mean Embeddings

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Moments and cumulants on $\mathbb{R} \ni X \sim \gamma$

- Moments $\mu(\gamma) := (\mu^{(i)}(\gamma))_{i \in \mathbb{N}}$:

$$\mu^{(i)}(\gamma) := \mathbb{E}(X^i) \in \mathbb{R}, \quad \mu^{(0)}(\gamma) := 1.$$

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- Cumulants $\kappa(\gamma) = \left(\kappa^{(i)}(\gamma) \right)_{i \in \mathbb{N}}$: from the **moment-generating function**

$$\sum_{i \in \mathbb{N}} \kappa^{(i)}(\gamma) \frac{\theta^i}{i!} = \log \left(\sum_{i \in \mathbb{N}} \mu^{(i)}(\gamma) \frac{\theta^i}{i!} \right).$$

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$\kappa^{(1)}(\gamma) = \mathbb{E}(X)$	mean
$\kappa^{(2)}(\gamma) = \mathbb{E}(X - \mathbb{E}X)^2$	variance
$\kappa^{(3)}(\gamma) = \mathbb{E}(X - \mathbb{E}X)^3$	3rd central moment
$\kappa^{(4)}(\gamma) = \mathbb{E}(X - \mathbb{E}X)^4 - 3 [\mathbb{E}(X - \mathbb{E}X)^2]^2$	
$\kappa^{(5)}(\gamma) = \mathbb{E}(X - \mathbb{E}X)^5 - 10\mathbb{E}(X - \mathbb{E}X)^3\mathbb{E}(X - \mathbb{E}X)^2$	

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...

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Question

What are the **weights** in front of the moments?

Unzipping cumulants on \mathbb{R} : the weights

m	elements of $\pi \in P(m)$	$ \pi $	c_π
1	{1}	1	1
2	{1,2}	1	1
	{1},{2}	2	-1
3	{1,2,3}	1	1
	{1,2}, {3}	2	-1
	{1,3}, {2}	2	-1
	{2,3}, {1}	2	-1
	{1}, {2}, {3}	3	2

with $P(m) :=$ all partitions of $[m]$, $c_\pi = (-1)^{|\pi|-1}(|\pi| - 1)!$

Motivation, i.e. one reason why one likes cumulants

Moment and cumulants on \mathbb{R}^d

Change $\mathbb{E}(\textcolor{teal}{X}^i) \in \mathbb{R}$ to $\mathbb{E}\left[X_1^{i_1} \cdots X_d^{i_d}\right] \in \mathbb{R}$ ($\mathbf{i} \in \mathbb{N}^d$). log, $P(m)$: ✓

Known theorem [Billingsley, 2012]

Let γ be a probability measure on a bounded subset of \mathbb{R}^d with cumulants $\kappa(\gamma)$ and let $(X_1, \dots, X_d) \sim \gamma$. Then

- ① $\gamma \mapsto \kappa(\gamma)$ is injective.
- ② X_1, \dots, X_d are independent $\Leftrightarrow \kappa^{\mathbf{i}}(\gamma) = 0$ for all $\mathbf{i} \in \mathbb{N}_+^d$.

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Motivation

- ① Various data types, nonlinear features: kernels.
- ② Linear: not even characteristic (see MMD and HSIC).
- ③ Computable estimators.

Idea

Lifting

$$(X_1, \dots, X_d) \in \times_{j=1}^d \mathcal{X}_j \rightarrow (\Phi_1(X_1), \dots, \Phi_d(X_d)) \in \times_{j=1}^d \mathcal{H}_{k_j}.$$

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Ingredients:

- ➊ Moments: swap out $\mathbb{E} [X_1^{i_1} \cdots X_d^{i_d}] \in \mathbb{R}$ to

$$\mathbb{E} [\Phi_1(X_1)]^{\otimes i_1} \otimes \cdots \otimes [\Phi_d(X_d)]^{\otimes i_d} \in \mathcal{H}_{k_1}^{\otimes i_1} \otimes \cdots \otimes \mathcal{H}_{k_d}^{\otimes i_d}.$$

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- ② From moments to cumulants:

- log on tensor algebras, or
- combinatorial description of cumulants (\leftarrow a bit simpler, but \Leftrightarrow).

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- ② From moments to cumulants:
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- ③ Computation: by the 'expected kernel trick' (V-statistics).

Kernel (generalization of $\mathbf{a}^\top \mathbf{b}$), RKHS

- Def-1 (feature space):

$$k(a, b) = \langle \Phi(a), \Phi(b) \rangle_{\mathcal{H}}.$$

- Def-2 (reproducing kernel):

$$k(\cdot, b) \in \mathcal{H}, \quad f(b) = \langle f, k(\cdot, b) \rangle_{\mathcal{H}}.$$

- Def-3 (Gram matrix): $\mathbf{G} = [k(x_i, x_j)]_{i,j=1}^n \in \mathbb{R}^{n \times n} \succeq 0$.

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Notes

- $k \stackrel{1:1}{\leftrightarrow} \mathcal{H}_k = \overline{\text{Span}}(k(\cdot, x) : x \in \mathcal{X})$: Fourier analysis, polynomials, splines, ...
- Examples: $k_p(\mathbf{x}, \mathbf{y}) = (\langle \mathbf{x}, \mathbf{y} \rangle + \gamma)^p$, $k_G(\mathbf{x}, \mathbf{y}) = e^{-\gamma \|\mathbf{x} - \mathbf{y}\|_2^2}$.
- Kernels exist on various domains!

Mean embedding

- Mean embedding (Bochner integral):

$$\mu_k(\mathbb{P}) := \int_{\mathcal{X}} \underbrace{k(\cdot, x)}_{\Phi(x) \in \mathcal{H}_k} d\mathbb{P}(x) \in \mathcal{H}_k.$$

Mean embedding, MMD

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$$\text{MMD}_k(\mathbb{P}, \mathbb{Q}) := \|\mu_k(\mathbb{P}) - \mu_k(\mathbb{Q})\|_{\mathcal{H}_k}.$$

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Clarification of what $\otimes_{j=1}^d k_j$ and $\otimes_{j=1}^d \mu_{k_j}(\mathbb{P}_j)$ are follows.

Tensor product: $\otimes_{j=1}^d a_j$

- If $\mathbf{a} \in \mathbb{R}^{n_1}$, $\mathbf{b} \in \mathbb{R}^{n_2}$:

$$\mathbb{R} \ni \mathbf{v}^\top (\mathbf{a} \mathbf{b}^\top) \mathbf{w} = (\mathbf{v}^\top \mathbf{a}) (\mathbf{b}^\top \mathbf{w}) = \langle \mathbf{a}, \mathbf{v} \rangle_{\mathbb{R}^{n_1}} \langle \mathbf{b}, \mathbf{w} \rangle_{\mathbb{R}^{n_2}},$$

$\mathbf{a} \otimes \mathbf{b} := \mathbf{a} \mathbf{b}^\top$ is an $\mathbb{R}^{n_1} \times \mathbb{R}^{n_2} \rightarrow \mathbb{R}$ bilinear form.

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- For $a \in \mathcal{H}_1$, $b \in \mathcal{H}_2$ Hilbert spaces, i.e. for $d = 2$:

$$(a \otimes b)(v, w) := \langle a, v \rangle_{\mathcal{H}_1} \langle b, w \rangle_{\mathcal{H}_2}.$$

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- For $d \geq 2$ and $a_j \in \mathcal{H}_j$,

$$\left(\otimes_{j=1}^d a_j \right) (b_1, \dots, b_d) := \prod_{j=1}^d \langle a_j, b_j \rangle_{\mathcal{H}_j}.$$

Tensor product: $\otimes_{j=1}^d \mathcal{H}_j$

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$\xrightarrow{\text{spec.}}$ The tensor product of RKHSs is an RKHS

$$\mathcal{H}_k = \otimes_{j=1}^d \mathcal{H}_{k_j},$$

$$k(x, x') := (\otimes_{j=1}^d k_j)(x, x') := \prod_{j=1}^d \underbrace{k_j(x_j, x'_j)}_{\text{coordinate-wise similarity}}.$$

Validness of MMD & HSIC, their estimation

Validness:

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Properties:

- ① Injectivity of μ_k on probability / finite signed measures, so universal \Rightarrow characteristic.

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Properties:

- ① Injectivity of μ_k on probability / finite signed measures, so universal \Rightarrow characteristic.
- ② Easy-to-estimate: expected kernel trick

$$\langle \mu_k(\mathbb{P}), \mu_k(\mathbb{Q}) \rangle_{\mathcal{H}_k} = \int_{\mathcal{X}} \int_{\mathcal{X}} k(x, y) d\mathbb{P}(x) d\mathbb{Q}(y).$$

Kernelized moments – towards kernelized cumulants

- From now:

- $X = (X_j)_{j=1}^d \in \times_{j=1}^d \mathcal{X}_j$, $X \sim \gamma$,
- kernels $k_j : \mathcal{X}_j \times \mathcal{X}_j \rightarrow \mathbb{R}$, $j \in [d]$,
- lifting $\Phi(X) = (\Phi_j(X_j))_{j=1}^d$ with $\Phi_j(x_j) := k_j(\cdot, x_j)$,
- RKHS $\mathcal{H}^{\otimes \mathbf{i}} := \mathcal{H}_{k_1}^{\otimes i_1} \otimes \cdots \otimes \mathcal{H}_{k_d}^{\otimes i_d}$ with kernel $k^{\otimes \mathbf{i}} := k_1^{\otimes i_1} \otimes \cdots \otimes k_d^{\otimes i_d}$,
and feature

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- Moment sequence:

$$\mu(\gamma) = \left(\mu^{\mathbf{i}}(\gamma) \right)_{\mathbf{i} \in \mathbb{N}^d}, \quad \mu^{\mathbf{i}}(\gamma) := \mathbb{E} [\Phi^{\otimes \mathbf{i}}(X)] \in \mathcal{H}^{\otimes \mathbf{i}}.$$

Kernelized cumulants: examples first, analogous to \mathbb{R}

- $d = 1, m \in [3]: X \sim \gamma,$

$$\kappa_k^{(1)}(\gamma) = \mathbb{E}[\Phi(X)]$$

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$$+ 2\mathbb{E}^{\otimes 3}[\Phi(X)].$$

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- $d = 2, m = 2$: $(X_1, X_2) \sim \gamma$,

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$$\kappa_{k_1, k_2}^{(2,0)}(\gamma) = \mathbb{E}[\Phi_1^{\otimes 2}(X_1)] - \mathbb{E}^{\otimes 2}[\Phi_1(X_1)],$$

$$\kappa_{k_1, k_2}^{(1,1)}(\gamma) = \mathbb{E}[\Phi_1(X_1) \otimes \Phi_2(X_2)] - \mathbb{E}[\Phi_1(X_1)] \otimes \mathbb{E}[\Phi_2(X_2)],$$

$$\kappa_{k_1, k_2}^{(0,2)}(\gamma) = \mathbb{E}[\Phi_2^{\otimes 2}(X_2)] - \mathbb{E}^{\otimes 2}[\Phi_2(X_2)].$$

Wanted: repetition and partitioning. Weights: as before (c_π).

Kernelized cumulants: $X \sim \gamma$ prob. measure on $\times_{j=1}^d \mathcal{X}_j$

- Repetition (diagonal measure): $\mathbf{i} \in \mathbb{N}^d$,

$$\gamma^{\mathbf{i}} := \text{Law}\left(\underbrace{X_1, \dots, X_1}_{i_1 \text{ times}}, \underbrace{X_2, \dots, X_2}_{i_2 \text{ times}}, \dots, \underbrace{X_d, \dots, X_d}_{i_d \text{ times}}\right).$$

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- Kernelized cumulants: $\mathbf{m} = \deg(\mathbf{i}) := \sum_{j=1}^d i_j \xrightarrow{\text{OK}} \gamma_\pi^{\mathbf{i}} = (\gamma^{\mathbf{i}})_\pi$,

$$\kappa_{k_1, \dots, k_d}(\gamma) := \left(\kappa_{k_1, \dots, k_d}^{\mathbf{i}}(\gamma) \right)_{\mathbf{i} \in \mathbb{N}^d},$$

$$\kappa_{k_1, \dots, k_d}^{\mathbf{i}}(\gamma) := \sum_{\pi \in P(\mathbf{m})} c_\pi \mathbb{E}_{\gamma_\pi^{\mathbf{i}}} k^{\otimes \mathbf{i}}(\cdot, (X_1, \dots, X_{\mathbf{m}})).$$

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⇒ expected kernel trick is applicable

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Point-separating k := injectivity of $\Phi \Leftarrow$ characteristic $k \Leftarrow$ universal k .

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Theorem

- Assume:

- γ, η : probability measures on $\times_{j=1}^d \mathcal{X}_j$,
- $(\mathcal{X}_j)_{j=1}^d$ are Polish spaces,
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- Then, $\gamma = \eta \Leftrightarrow \kappa_{k_1, \dots, k_d}(\gamma) = \kappa_{k_1, \dots, k_d}(\eta)$, and
$$\begin{aligned} d^{\mathbf{i}}(\gamma, \eta) &:= \|\kappa_{k_1, \dots, k_d}^{\mathbf{i}}(\gamma) - \kappa_{k_1, \dots, k_d}^{\mathbf{i}}(\eta)\|_{\mathcal{H}^{\otimes \mathbf{i}}}^2 \\ &= \sum_{\pi, \tau \in P(m)} c_\pi c_\tau \left[\mathbb{E}_{\gamma_\pi^{\mathbf{i}} \otimes \gamma_\tau^{\mathbf{i}}} k^{\otimes \mathbf{i}}((X_1, \dots, X_m), (Y_1, \dots, Y_m)) \right. \\ &\quad + \mathbb{E}_{\eta_\pi^{\mathbf{i}} \otimes \eta_\tau^{\mathbf{i}}} k^{\otimes \mathbf{i}}((X_1, \dots, X_m), (Y_1, \dots, Y_m)) \\ &\quad \left. - 2\mathbb{E}_{\gamma_\pi^{\mathbf{i}} \otimes \eta_\tau^{\mathbf{i}}} k^{\otimes \mathbf{i}}((X_1, \dots, X_m), (Y_1, \dots, Y_m)) \right]. \end{aligned}$$

Cumulants characterize independence

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$$\|\kappa_{k_1, \dots, k_d}^{\mathbf{i}}(\gamma)\|_{\mathcal{H}^{\otimes \mathbf{i}}}^2 = \sum_{\pi, \tau \in P(m)} c_\pi c_\tau \mathbb{E}_{\gamma_\pi^{\mathbf{i}} \otimes \gamma_\tau^{\mathbf{i}}} k^{\otimes \mathbf{i}}((X_j)_{j=1}^m, (Y_j)_{j=1}^m),$$

where $m = \deg(\mathbf{i})$.

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where $m = \deg(\mathbf{i})$.

Estimation in both cases

$\mathbb{E} k^{\otimes \mathbf{i}}((X_1, \dots, X_m), (Y_1, \dots, Y_m)) \Rightarrow$ V-statistics ✓

Distance between kernel variance embeddings

- By our theorem: if $\gamma = \eta$, then $d^{(2)}(\gamma, \eta) = 0$.
- V-statistic estimator of $d^{(2)}(\gamma, \eta)$:

$$\frac{1}{N^2} \text{Tr}[(\mathbf{K}_x \mathbf{J}_N)^2] + \frac{1}{M^2} \text{Tr}[(\mathbf{K}_y \mathbf{J}_M)^2] - \frac{2}{NM} \text{Tr}[\mathbf{K}_{xy} \mathbf{J}_M \mathbf{K}_{xy}^\top \mathbf{J}_N],$$

with $(x_n)_{n=1}^N \stackrel{\text{i.i.d.}}{\sim} \gamma$, $(y_m)_{m=1}^M \stackrel{\text{i.i.d.}}{\sim} \eta$, $\mathbf{K}_x = [k(x_i, x_j)]_{i,j=1}^N$,
 $\mathbf{K}_y = [k(y_i, y_j)]_{i,j=1}^M$, $\mathbf{K}_{x,y} = [k(x_i, y_j)]_{i,j=1}^{N,M}$, $\mathbf{J}_n = \mathbf{I}_n - \mathbf{H}_n = \mathbf{I}_n - \frac{1}{n} \mathbf{1}_n \mathbf{1}_n^\top$.

Distance between kernel variance/skewness embeddings

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- V-statistic estimator of $d^{(2)}(\gamma, \eta)$:

$$\frac{1}{N^2} \text{Tr}[(\mathbf{K}_x \mathbf{J}_N)^2] + \frac{1}{M^2} \text{Tr}[(\mathbf{K}_y \mathbf{J}_M)^2] - \frac{2}{NM} \text{Tr}[\mathbf{K}_{xy} \mathbf{J}_M \mathbf{K}_{xy}^\top \mathbf{J}_N],$$

with $(x_n)_{n=1}^N \stackrel{\text{i.i.d.}}{\sim} \gamma$, $(y_m)_{m=1}^M \stackrel{\text{i.i.d.}}{\sim} \eta$, $\mathbf{K}_x = [k(x_i, x_j)]_{i,j=1}^N$,
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Time complexity

Quadratic as MMD.

- $d^{(3)}(\gamma, \eta)$: similarly; quadratic time.

Cross-skewness independence criterion (CSIC)

- By our theorem: if $\gamma = \gamma|x_1 \otimes \gamma|x_2$, then $\kappa_{k,\ell}^{(2,1)}(\gamma) = 0$ and $\kappa_{k,\ell}^{(1,2)}(\gamma) = 0$.
- V-statistic estimator of $\|\kappa_{k,\ell}^{(1,2)}(\gamma)\|_{\mathcal{H}_k^{\otimes 1} \otimes \mathcal{H}_\ell^{\otimes 2}}^2$:

$$\begin{aligned} & \frac{1}{N^2} \left\langle \mathbf{K} \circ \mathbf{K} \circ \mathbf{L} - 4\mathbf{K} \circ \mathbf{K}\mathbf{H} \circ \mathbf{L} - 2\mathbf{K} \circ \mathbf{K} \circ \mathbf{L}\mathbf{H} + 4\mathbf{K}\mathbf{H} \circ \mathbf{K} \circ \mathbf{L}\mathbf{H} \right. \\ & \quad \left. + 2\mathbf{K} \circ \mathbf{L} \left\langle \frac{\mathbf{K}}{N^2} \right\rangle + 2\mathbf{K}\mathbf{H} \circ \mathbf{H}\mathbf{K} \circ \mathbf{L} + 4\mathbf{K} \circ \mathbf{H}\mathbf{K} \circ \mathbf{L}\mathbf{H} + \mathbf{K} \circ \mathbf{K} \left\langle \frac{\mathbf{L}}{N^2} \right\rangle \right. \\ & \quad \left. - 8\mathbf{K} \circ \mathbf{L}\mathbf{H} \left\langle \frac{\mathbf{K}}{N^2} \right\rangle - 4\mathbf{K} \circ \mathbf{H}\mathbf{K} \left\langle \frac{\mathbf{L}}{N^2} \right\rangle + 4 \left\langle \frac{\mathbf{K}}{N^2} \right\rangle^2 \mathbf{L} \right\rangle, \end{aligned}$$

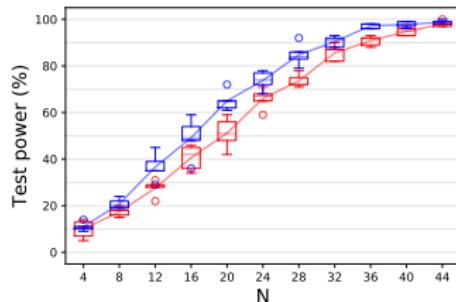
with kernels $k : \mathcal{X}_1^2 \rightarrow \mathbb{R}$, $\ell : \mathcal{X}_2^2 \rightarrow \mathbb{R}$, $\mathbf{K} := \mathbf{K}_x$, $\mathbf{L} := \mathbf{L}_y$, $\langle \mathbf{A} \rangle := \sum_{i,j} A_{i,j}$.

- Time complexity: quadratic.

Numerical illustration: improved power

- Seoul bicycle rental data:

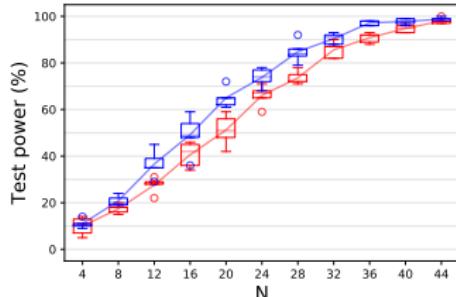
- two-sample test (MMD, $d^{(2)}$): winter vs fall, $d = 11$,



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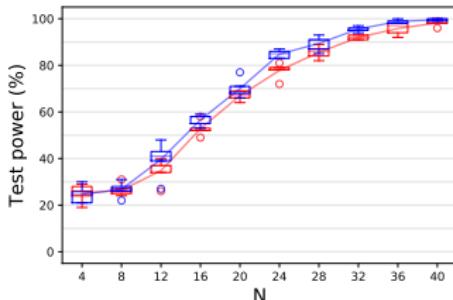
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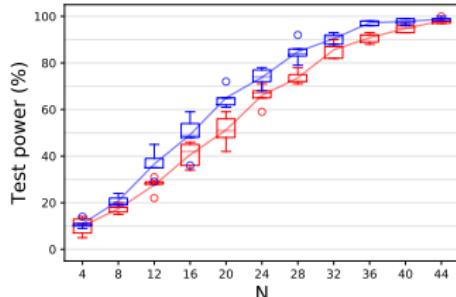
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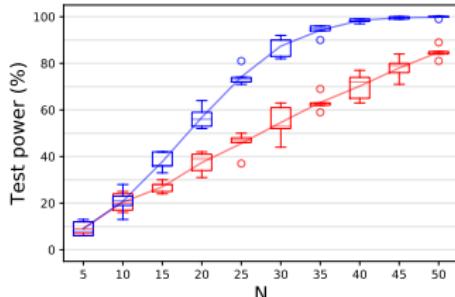
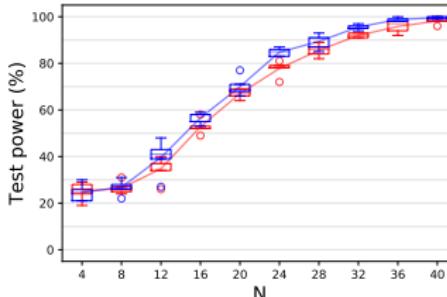
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- two-sample test (**MMD**, $d^{(3)}$): slow vs fast moving traffic, $d = 16$; r.h.s.



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Appendix

- Bell numbers
- Characteristic kernels
- Universal kernels
- Moments and cumulants on \mathbb{R}^d
- Estimator for $d^{(3)}(\gamma, \eta)$

Bell numbers

- $B(m) :=$ number of elements in $P(m)$.
- $B_0 = B_1 = 1, B_2 = 2, B_3 = 5, B_4 = 15, B_5 = 52, B_6 = 203,$
 $B_7 = 877, B_8 = 4140, \dots$

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 $B_7 = 877, B_8 = 4140, \dots$
- Recursion:

$$B_{m+1} = |P(m+1)| = \sum_{k=0}^m \binom{m}{k} B_k.$$

Bell numbers – continued

- Easy computation by the Bell triangle

1					
1	2				
2	3	5			
5	7	10	15		
15	20	27	37	52	
52	...				

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- Asymptotics:

$$\frac{\ln B_n}{n} = \ln n - \ln \ln n - 1 + \frac{\ln \ln n}{\ln n} + \frac{1}{\ln n} + \frac{1}{2} \left(\frac{\ln \ln n}{\ln n} \right)^2 + \mathcal{O} \left(\frac{\ln \ln n}{\ln^2 n} \right)$$

as $n \rightarrow \infty$.

Contents

Description of characteristic kernels on \mathbb{R}^d

For continuous bounded shift-invariant kernels on \mathbb{R}^d :

$$k(\mathbf{x}, \mathbf{x}') = k_0(\mathbf{x} - \mathbf{x}') \stackrel{(*)}{=} \int_{\mathbb{R}^d} e^{-i\langle \mathbf{x}-\mathbf{x}', \omega \rangle} d\Lambda(\omega)$$

(*): Bochner's theorem.

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Theorem ([Sriperumbudur et al., 2010])

k is characteristic iff. $\text{supp}(\Lambda) = \mathbb{R}^d$.

Examples on \mathbb{R} ; similarly \mathbb{R}^d [Sriperumbudur et al., 2010]

For Poisson kernel: $\sigma \in (0, 1)$.

kernel name	k_0	$\hat{k}_0(\omega)$	$supp(\hat{k}_0)$
Gaussian	$e^{-\frac{x^2}{2\sigma^2}}$	$\sigma e^{-\frac{\sigma^2\omega^2}{2}}$	\mathbb{R}
Laplacian	$e^{-\sigma x }$	$\sqrt{\frac{2}{\pi}} \frac{\sigma}{\sigma^2 + \omega^2}$	\mathbb{R}
B_{2n+1} -spline	$*^{2n+2} \chi_{[-\frac{1}{2}, \frac{1}{2}]}(x)$	$\frac{4^{n+1}}{\sqrt{2\pi}} \frac{\sin^{2n+2}\left(\frac{\omega}{2}\right)}{\omega^{2n+2}}$	\mathbb{R}
Sinc	$\frac{\sin(\sigma x)}{x}$	$\sqrt{\frac{\pi}{2}} \chi_{[-\sigma, \sigma]}(\omega)$	$[-\sigma, \sigma]$
Poisson	$\frac{1 - \sigma^2}{\sigma^2 - 2\sigma \cos(x) + 1}$	$\sqrt{2\pi} \sum_{j=-\infty}^{\infty} \sigma^{ j } \delta(\omega - j)$	\mathbb{Z}
Dirichlet	$\frac{\sin\left(\frac{(2n+1)x}{2}\right)}{\sin\left(\frac{x}{2}\right)}$	$\sqrt{2\pi} \sum_{j=-\infty}^{\infty} \delta(\omega - j)$	$\{0, \pm 1, \pm 2, \dots, \pm n\}$
Fejér	$\frac{1}{n+1} \frac{\sin^2\left(\frac{(n+1)x}{2}\right)}{\sin^2\left(\frac{x}{2}\right)}$	$\sqrt{2\pi} \sum_{j=-n}^n \left(1 - \frac{ j }{n+1}\right) \delta(\omega - j)$	$\{0, \pm 1, \pm 2, \dots, \pm n\}$
Cosine	$\cos(\sigma x)$	$\sqrt{\frac{\pi}{2}} [\delta(\omega - \sigma) + \delta(\omega + \sigma)]$	$\{-\sigma, \sigma\}$

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Cosine	$\cos(\sigma x)$	$\sqrt{\frac{\pi}{2}} [\delta(\omega - \sigma) + \delta(\omega + \sigma)]$	$\{-\sigma, \sigma\}$

For $x \in \mathbb{R}^d$: $k_0(x) = \prod_{j=1}^d k_0(x_j)$, $\widehat{k}_0(\omega) = \prod_{j=1}^d \widehat{k}_0(\omega_j)$.

Properties of universal kernels

[Steinwart, 2001, Steinwart and Christmann, 2008]

If k is universal, then

- $k(x, x) > 0$ for all $x \in \mathcal{X}$.

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- Every restriction of k to an $\mathcal{X}' \subseteq \mathcal{X}$ compact set is universal.
- $\Phi(x) = k(\cdot, x)$ is injective, i.e.

$$\rho_k(x, y) = \|\Phi(x) - \Phi(y)\|_{\mathcal{H}_k}$$

is a metric.

Properties of universal kernels

[Steinwart, 2001, Steinwart and Christmann, 2008]

If k is universal, then

- $k(x, x) > 0$ for all $x \in \mathcal{X}$.
- Every restriction of k to an $\mathcal{X}' \subseteq \mathcal{X}$ compact set is universal.
- $\Phi(x) = k(\cdot, x)$ is injective, i.e.

$$\rho_k(x, y) = \|\Phi(x) - \Phi(y)\|_{\mathcal{H}_k}$$

is a metric.

- The normalized kernel (like corr)

$$\tilde{k}(x, y) := \frac{k(x, y)}{\sqrt{k(x, x)k(y, y)}}$$

is universal.

Universal Taylor kernels

[Steinwart, 2001, Steinwart and Christmann, 2008]

- For an $C^\infty \ni f : (-r, r) \rightarrow \mathbb{R}$

$$f(t) = \sum_{n=0}^{\infty} \textcolor{blue}{a_n} t^n \quad t \in (-r, r), \quad r \in (0, \infty].$$

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- If $a_n > 0 \ \forall n$, then

$$k(\mathbf{x}, \mathbf{y}) = f(\langle \mathbf{x}, \mathbf{y} \rangle)$$

is universal on $\mathcal{X} := \left\{ \mathbf{x} \in \mathbb{R}^d : \|\mathbf{x}\|_2 \leq \sqrt{r} \right\}$.

Universal kernels on compact subsets of \mathbb{R}^d , $\alpha > 0$

- $k(\mathbf{x}, \mathbf{y}) = e^{\alpha \langle \mathbf{x}, \mathbf{y} \rangle}$: previous result with $f(t) = e^{\alpha t} \Rightarrow a_n = \frac{\alpha^n}{n!}$.

Universal kernels on compact subsets of \mathbb{R}^d , $\alpha > 0$

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- $k(\mathbf{x}, \mathbf{y}) = e^{-\alpha \|\mathbf{x} - \mathbf{y}\|_2^2}$: exp. kernel & normalization.

Universal kernels on compact subsets of \mathbb{R}^d , $\alpha > 0$

- $k(\mathbf{x}, \mathbf{y}) = (1 - \langle \mathbf{x}, \mathbf{y} \rangle)^{-\alpha}$ binomial kernel
 - on \mathcal{X} compact $\subset \{\mathbf{x} \in \mathbb{R}^d : \|\mathbf{x}\|_2 < 1\}$.
 - $f(t) = (1 - t)^{-\alpha} = \underbrace{\sum_{n=0}^{\infty} \binom{-\alpha}{n} (-1)^n t^n}_{>0} \quad (|t| < 1),$
- where $\binom{b}{n} = \sum_{i=1}^n \frac{b-i+1}{i}$.

Contents

Moments and cumulants on $\mathbb{R}^d \ni X \sim \gamma$, $\mathbf{i} \in \mathbb{N}^d$

	$d = 1$	$d \geq 1$
moment sequence	$\mu(\gamma) := \left(\mu^{(i)}(\gamma) \right)_{i \in \mathbb{N}}$	$\mu(\gamma) := \left(\mu^{\mathbf{i}}(\gamma) \right)_{\mathbf{i} \in \mathbb{N}^d}$
moments	$\mu^{(i)}(\gamma) := \mathbb{E} (\textcolor{green}{X}^{\mathbf{i}}) \in \mathbb{R}$	$\mu^{\mathbf{i}}(\gamma) := \mathbb{E} [X_1^{i_1} \cdots X_d^{i_d}] \in \mathbb{R}$

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where $\deg(\mathbf{i}) := i_1 + \cdots + i_d$, $\mu^0(\gamma) = 1$

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and cumulants $\kappa(\gamma) = (\kappa^{\mathbf{i}}(\gamma))_{\mathbf{i} \in \mathbb{N}^d}$

$$\sum_{\mathbf{i} \in \mathbb{N}^d} \kappa^{\mathbf{i}}(\gamma) \frac{\theta^{\mathbf{i}}}{\mathbf{i}!} = \log \left(\sum_{\mathbf{i} \in \mathbb{N}^d} \mu^{\mathbf{i}}(\gamma) \frac{\theta^{\mathbf{i}}}{\mathbf{i}!} \right), \quad \theta \in \mathbb{R}^d,$$

where $\deg(\mathbf{i}) := i_1 + \cdots + i_d$, $\mu^0(\gamma) = 1$, $\mathbf{i}! = i_1! \cdots i_d!$, $\theta^{\mathbf{i}} = \theta_1^{i_1} \cdots \theta_d^{i_d}$.

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Estimator for $d^{(3)}(\gamma, \eta) = \|\kappa_k^{(3)}(\gamma) - \kappa_k^{(3)}(\eta)\|_{\mathcal{H}_k^{\otimes 3}}^2$, $N = M$

$$d^{(3)}(\gamma, \eta) = \|\kappa_k^{(3)}(\gamma)\|_{\mathcal{H}_k^{\otimes 3}}^2 + \|\kappa_k^{(3)}(\eta)\|_{\mathcal{H}_k^{\otimes 3}}^2 - 2\langle \kappa_k^{(3)}(\gamma), \kappa_k^{(3)}(\eta) \rangle_{\mathcal{H}_k^{\otimes 3}}$$

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$$\begin{aligned} \langle \kappa_k^{(3)}(\gamma), \kappa_k^{(3)}(\eta) \rangle_{\mathcal{H}_k^{\otimes 3}} &\approx \frac{1}{N^2} \left\langle \mathbf{K}_{xy} \circ \mathbf{K}_{xy} \circ \mathbf{K}_{xy} - 3\mathbf{K}_{xy} \circ \mathbf{K}_{xy} \circ \mathbf{H}\mathbf{K}_{xy} \right. \\ &\quad - 3\mathbf{K}_{xy} \circ \mathbf{K}_{xy} \circ \mathbf{K}_{xy}\mathbf{H} + 6\mathbf{K}_{xy} \circ \mathbf{K}_{xy}\mathbf{H} \circ \mathbf{H}\mathbf{K}_{xy} \\ &\quad + 3\mathbf{K}_{xy} \circ \mathbf{K}_{xy} \left\langle \frac{\mathbf{K}_{xy}}{N^2} \right\rangle + 2\mathbf{K}_{xy} \circ \mathbf{H}\mathbf{K}_{xy} \circ \mathbf{H}\mathbf{K}_{xy} \\ &\quad + 2\mathbf{K}_{xy} \circ \mathbf{K}_{xy}\mathbf{H} \circ \mathbf{K}_{xy}\mathbf{H} - 6\mathbf{K}_{xy} \circ \mathbf{K}_{xy}\mathbf{H} \left\langle \frac{\mathbf{K}_{xy}}{N^2} \right\rangle \\ &\quad \left. - 6\mathbf{K}_{xy} \circ \mathbf{H}\mathbf{K}_{xy} \left\langle \frac{\mathbf{K}_{xy}}{N^2} \right\rangle + 4 \left\langle \frac{\mathbf{K}}{N^2} \right\rangle^2 \mathbf{K}_{xy} \right\rangle. \end{aligned}$$

Note: Matrix multiplication takes precedence over the Hadamard one.

Estimator for $d^{(3)}(\gamma, \eta)$ – continued

$$\begin{aligned}\|\kappa_k^{(3)}(\gamma)\|_{\mathcal{H}_k^{\otimes 3}}^2 &\approx \frac{1}{N^2} \left\langle \mathbf{K}_x \circ \mathbf{K}_x \circ \mathbf{K}_x - 6\mathbf{K}_x \circ \mathbf{K}_x \mathbf{H} \circ \mathbf{K}_x \right. \\ &\quad + 4\mathbf{K}_x \mathbf{H} \circ \mathbf{K}_x \circ \mathbf{K}_x \mathbf{H} + 3\mathbf{K}_x \circ \mathbf{K}_x \left\langle \frac{\mathbf{K}_x}{N^2} \right\rangle \\ &\quad + 6\mathbf{K}_x \mathbf{H} \circ \mathbf{H} \mathbf{K}_x \circ \mathbf{K}_x - 12\mathbf{K}_x \circ \mathbf{H} \mathbf{K}_x \left\langle \frac{\mathbf{K}_x}{N^2} \right\rangle \\ &\quad \left. + 4 \left\langle \frac{\mathbf{K}_x}{N^2} \right\rangle^2 \mathbf{K}_x \right\rangle.\end{aligned}$$

$\|\kappa_k^{(3)}(\eta)\|_{\mathcal{H}_k^{\otimes 3}}^2$: similarly (change \mathbf{K}_x to \mathbf{K}_y).

[Contents](#), [d²\(γ, η\) estimation](#)

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