# Cross-fitted instrument: a blueprint for one-sample Mendelian Randomization

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## Instrumental variable



# Mendelian Randomisation: using genotype as an instrument

- First idea in by Katan, APOUPOPROTEIN E ISOFORMS, SERUM CHOLESTEROL, AND CANCER, *Lancet*, 1986
- R Gray, K Wheatley, How to avoid bias when comparing bone marrow transplantation with chemotherapy, *Bone Marrow Transplantation*, 1991
- Davey Smith, Mendelian Randomization for Strengthening Causal Inference in Observational Studies: Application to Gene × Environment Interactions, *Perspectives on Psychological Science*, 2010

# Random inheritance of genotype



Figure from Bates *et al.* Causal Inference in Genetic Trio Studies, *PNAS*, 2021

## Using genotype as an instrument



### Using genotype as instruments



## Selection of the instruments: GWAS



### The basic strategy: one sample MR



### The basic strategy: one sample MR



2) For each variant selected: two stage least square

# Endogeneity/weak instrument bias

$$Y = \beta_{X \to Y} X + H + U, \qquad \mathbb{E}[U|\Pi, X] = 0$$
(1)  
$$X = Z\Pi + H + V, \qquad \mathbb{E}[V|X] = 0$$
(2)

- $\beta_{X \to Y}$  is the effect of X on Y
- $\Pi$  is the vector of regression coefficients for the instruments
- U and V are two correlated error terms
- H hidden confounder

# Endogeneity/weak instrument bias

Nagar, The bias and moment matrix of the general k-class estimators of the parameters in simultaneous equations. *Econometrica* 1959

Bias tsls 
$$\approx \frac{\sigma_{U,V}}{\mathbb{E}(F)\sigma_U^2}$$
 (3)

- $\sigma_{U,V}$  covariance of the error terms in the first- and second-stage regression models
- $F \approx$  strength of the instruments, sample size

#### Reducing the bias from tsls

- Increase sample size
- Ind stronger instruments: SNPs have small effect size
- Set  $\sigma_{U,V}$  to 0

## The two-sample MR



# The two-sample MR: source of bias

#### Sample overlap

• Burgess and colleagues (2016) showed that if sample 1 and sample 2 are overlapping, endogeneity bias has to be expected. (Mounier and Kutalik, Correction for sample overlap, winner's curse and weak instrument bias in two-sample Mendelian Randomization, *BiorXiv*, March 28 2021)

#### Population heterogeneity

• The effect of a SNP can vary from a population to another (due to change in minor allele frequency). A SNP could be causal for the exposure (sample 1) but could be constant within another population.

# One sample MR

#### Pros

Homogeneous population

Fast

#### Cons

- Endogeneity bias/winner curse
- Overconfident confidence interval



# Two-sample MR

#### Pros

- Less prone to edogeneity bias
- Use of summary statistics available online

#### Cons

- Potentially unfeasible for rare or expensive phenotype
- Potentially slow
- Severe waste of data



# Getting the best of both worlds

#### Endogeneity free one sample MR

- Propose an approach that use only one sample and that has no endogeneity bias/winner's curse
- We developed the concept of cross-fitted instrument/cross-fitted instruments (CFI/CFIs)

### Build on

- Double Machine learning by Chernozhukov *et al., The Econometrics Journal*, 2017
- Older approaches such as Split sample IV or Jackknifed IV from Angrist and Krueger, 1995 and Angrist, Krueger and Imbens, 1999

CFI: middle ground between Split sample IV and Jackknifed IV

## Cross-fitted instrument

- 2-fold cross-fitted instruments
- k-fold cross-fitted instrument/instruments
- CFMR1 and CFMR2

# Sample splitting



## Selection of the instruments



### Two stage least squares



## Average



# Sample splitting





## Select instruments using samples 1 and 2



# Predict X in sample 3 using estimates from samples 1 and 2 $% \left( {\left[ {{X_{\rm{B}}} \right]_{\rm{B}}} \right)_{\rm{B}} \right)$



## Write the IV vector



## Select instruments using samples 2 and 3



# Predict X in sample 1 using estimates from samples 2 and 3 $% \left( {\left[ {{X_{\rm{B}}} \right]_{\rm{B}}} \right)$



## Select instruments using samples 1 and 3



# Predict X in sample 2 using estimates from samples 1 and 3 $% \left( {\left[ {{X_{\rm{B}}} \right]_{\rm{B}}} \right)$



## CFMR1 and CFMR2



# Simulations and application

- Endogeneity bias
- power of CFMR vs two-sample MR
- Estimating the effect of pre-pregnancy maternal BMI on child birth weight

# Bias in one sample MR

• We consider a set of 300 independent variants ( $V_1, ..., V_{300}$ )

• 
$$X = \sum_{l=1}^{5} \pi V_l + 40 \times h + v$$

- Y = 0.8X + h + u
- h is a hidden confounder generated from a N(0,2) distribution, and v and u are two correlated error terms generated from a bivariate normal distribution.

$$\begin{pmatrix} U \\ V \end{pmatrix} \sim N \begin{bmatrix} \begin{pmatrix} 0 \\ 0 \end{pmatrix}, & \begin{pmatrix} 1 & 0.9 \\ 0.9 & 1 \end{bmatrix} \end{bmatrix}$$

We consider the following two scenarios:

- **(**) where the variants explain 10% of the variance of X
- ② where the variants explain 20% of the variance of X

# Bias in one sample MR

#### For each simulated dataset

- We applied 10-fold CFMR1 using a LASSO-based IV.
- We also build a predictor of X using LASSO on the entire dataset. We then used the prediction on the entire data as an instrument. We refer to 'one-sample MR estimates' when we estimate the effect of X on Y

#### Nota bene:

• In our manuscript we show that CFMR remains conservative even when using instrument that explain only 0.001% of the exposure variance.



## Power comparison

• We consider a set of 300 independent variants ( $V_1, ..., V_{300}$ )

• 
$$X = \sum_{l=1}^{5} \pi V_l + h + v$$

- $Y = \theta_0 X + h + u$
- h is a hidden confounder generated from a N(0,2) distribution, and v and u are two correlated error terms generated from a bivariate normal distribution.

$$\begin{pmatrix} U \\ V \end{pmatrix} \sim N \begin{bmatrix} \begin{pmatrix} 0 \\ 0 \end{pmatrix}, & \begin{pmatrix} 1 & 0.2 \\ 0.2 & 1 \end{bmatrix} \end{bmatrix}$$

where the variants explain 20% of the variance of X Comparison with theoretical power of two-sample MR from Deng *et al.*, *Genetic Epidemiology*, 2020

# Power: thick lines from Deng et al., Genetic Epidemiology, 2020



# Estimating the effect of pre-pregnancy maternal BMI on childbirth weight

- We applied a 10 fold CFMR1 to a dataset comprising mother-child duos from the Norwegian Mother, father, and Child Cohort Study (MoBa), to re-examine the well-established effect of maternal pre-pregnancy BMI on offspring's birth weight (Tyrrel *et al., JAMA*, 2016).
- 26,896 complete mother-child duos with genotype and phenotype data remained for the current analyses.
- 10 separate GWASes of pre-pregnancy BMI performed, with each GWAS encompassing 24, 210 randomly selected mothers.

# Polygenic score for maternal BMI (p-value $10^{-3}$ )



# Polygenic score for maternal BMI (p-value $10^{-5}$ )



# CFMR estimates using different p-value threshold



 $\label{eq:cross-fitted instrument: a blueprint for one-sample Mendelian Randomization \\ \ensuremath{\mathsf{Application}}$ 

### Thank you for listening.

#### Joint work with:

- Jon Bohlin,
- Stephen Burgess,
- Christian Page,
- Astanand Jugessur

- Cross-fitted instrument: a blueprint for one-sample Mendelian Randomization, BiorXiv, 2021 (under review)
- https://github.com/williamdenault/CFMR