

MI and IPW methods for addressing both intermittent and monotone missingness in comparative effectiveness studies with timevarying confounding

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• Rationale

- Motivating example
- Simulation study

• Preliminary findings and next steps



- **Routine data** increasingly used to establish the effectiveness and cost-effectiveness of health interventions
- **NICE** and other agencies now routinely use these data to supplement trial evidence (when this is limited/inexistent)
- One such area is the evaluation of treatment strategies sustained over time, i.e. **time-varying treatments**
- Central feature in these studies is that treatment status is determined at different points over time
- Patient's progression typically influences future treatments and outcomes, and is itself affected by previous treatments

Rationale

- In addition to the confounding a recurrent issue in these studies is **missing data**
- Inverse probability weighting (IPW) tends to be most suitable to handle monotone missing data but it is less practical to handle intermittent missing data. Multiple Imputation (MI) known for its flexibility to address intermittent patterns
- A few studies have **compared MI and IPW** for addressing data in studies with time-varying confounding
 - Moodie et al 2008
 - Vourli and Touloumi 2014
 - Liu et al 2019
 - Leyrat et al 2021
- Focused mostly on missing confounders and one type of missingness (monotone)

Aim: To contrast IPW and MI methods for handling both monotone and intermittent missing data in both outcomes and confounders

Objectives:

- To illustrate how to combine MI with IPW-based marginal structural models (MSMs)
- To assess the performance of MI and IPW in settings with both monotone and intermittent missingness:
 - Method 1: use IPW for both monotone and intermittent
 - **Method 2**: use IPW for monotone, and MI for intermittent
- To illustrate the methods in an evaluation of biologic drugs for patients with severe rheumatoid arthritis

Rheumatoid arthritis (RA) study

- RCTs showed some **biologics** (e.g. Etanercept) effectively slow RA progression and improves patients' HRQL over short-term
- **No randomised evidence** on the sustained effect of these biologics over long term (key parameter for CER NICE decisions)
- We use **US National Data Bank** for rheumatic diseases and estimate 5-year effect of Etanercept vs other biologics on EQ-5D
- Restricted sample (N=13,002) to patients with severe RA who failed to respond to first-line treatment – biologic initiators
- Data are collected biannually as part of a clinical visit and also using patient-reported questionnaire

Missing data

id	time	At	Xt	Yt
1	1	A11	X11	•
1	2	A12	•	Y12
1	3	A13		•
1	4	A14	X14	Y14
2	1	A21	X21	Y21
2	2			
2	3			
2	4	A24	X24	Y24
3	1			
3	2	A32	X32	Y32
3	3	A33	X33	
3	4	A34	X34	•
4	1	A41	X41	Y41
4	2	A42	X42	Y42
4	3	A43	X43	Y43
4	4			

TOTAL ~25% missing

Intermittent missing X and Y (5%)

Interval Censoring (10%)

Left censoring (0.1%)

Monotone missing (4%)

Right censoring (6%)

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Time-varying confounding



- A Treatment (1 if Etanercept, 0 otherwise)
- X Time-varying confounder (HAQ disability index)
- Y Outcome (EQ-5D)

Simplified DAG: e.g. - no direct long-term effects of A on Y or X - no baseline variables



Inverse probability weighting (IPW or IPTW)

- IPW re-weights individuals according to the probability of being assigned to treatment (conditional on observed confounders)
- We used the recommended stabilised weights (Daniel et al 2012):

$$SW_{t} = \prod_{t=1}^{T} \frac{P(A_{t} | \bar{A}_{t-1}, B)}{P(A_{t} | \bar{A}_{t-1}, \bar{Y}_{t-1}, \bar{X}_{t}, B)}$$

- The relationship between treatment and outcome can then be modelled using a MSM on re-weighted sample:

$$(Y^{\bar{a}_T}) = \beta_0 + \sum_{t=1}^T \beta_t a_t$$



Method 1 – use IPW for both monotone and intermittent patterns

$$SC_{t} = \prod_{t=1}^{T} \frac{P(R_{t} = 1 | R_{t-1} = 1, B)}{P(R_{t} = 1 | \bar{A}_{t-1}, \bar{Y}_{t-1}, \bar{X}_{t-1}, R_{t-1} = 1, B)}$$

- Applied separately to monotone and intermittent missing data
- Discarded individuals when the treatment was missing at any point in time, had all outcome observations missing, or left censoring (to make it more comparable with simulations)

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$$W_t = SW_t * SCm_t * SCi_t$$

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Method 2 – use IPW for monotone and MI for intermittent missing data

For the IPW: $W_t = SW_t * SCm_t$

For the MI:

- Used chained equations and M = 10

-We followed Leyrat et al 2019's recommendations about combining MI with IPW weighting:

- Included outcome in the imputation model when imputing the confounders
- Applied Rubin's rules to treatment effects from MSM, rather than PS weights

RA study results





 Etanercept seems slightly more effective than other biologics in the first 2 years (as per trial evidence) but effect lost over time,



Data generating process (in line with DAG above)

- 2 time-constant continuous and binary covariates (Ba, Bs)
- 1 time-varying continuous confounder (X)
- 1 time-varying binary treatment indicator (T)
- 1 continuous outcome (Y)

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\begin{array}{l} \textit{Baseline (Time 0)} \\ \textit{Ba} = 14 + 76 * \textit{beta}(4.23, 2.7) \\ \textit{Bs} = 1(\textit{unif}(0, 1) < 0.06) \\ \textit{T}_0 = 0 \\ \textit{X}_0 = \textit{norm}(1.09 + s * 0.1 + a * 0.00001, 0.714) \\ \textit{Y}_0 = \textit{norm}(0.842 - \textit{X}_0 * 0.101 + s * 0.062 + 0.0001 * a, 0.152) \end{array}
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Subsequent periods

$$\begin{split} X_t = norm(1.62 + 0.38 * X_{t-1} + 0.7 * T_{t-1} - 1.32 * Y_{t-1} + 0.016 * \\ Bs + 0.001 * Ba, 0.716) \end{split}$$

$$T_{t} = 1(unif(0, 1))$$

$$< expit(-2.2 + 0.4 * X_{t} + 1.8 * T_{t-1} + 0.31 * Y_{t-1} + 0.041 * Bs + 0.025 * Ba))$$

 Y_t

 $= norm(0.32 - 0.036 * X_t + 0.05 * T_t + 0.63 * Y_{t-1} - 0.01 * s + 0.0001 * a, 0.144)$

Simulations – Missingness mechanism

Monotone/censored data ($C_0 = 0$) - MAR

$$C_t = 1(unif(0, 1))$$

< $(expit(-\beta_c - 2.1 * T_t + 0.25 * Bs + 0.025 * Ba) + C_{t-1} - \gamma_c)$

Intermittent missing outcome - MAR

$$Mo_{t} = 1(unif(0, 1))$$

< $(expit(-\beta_{Mo} - 1.1 * T_{t} + 0.8 * Bs + 0.045 * Ba) + C_{t-1} - \gamma_{Mo})$

Intermittent missing confounder - MAR

$$Mc_{t} = 1(unif(0, 1))$$

< $(expit(-\beta_{Mc} - 1.4 * T_{t} + 0.7 * Bs + 0.055 * Ba) + C_{t-1} - \gamma_{Mc})$

12 SCENARIOS

- Missing outcome, confounder, or both, plus usual censoring
- % missingness (low vs high, i.e. 20% vs 40%)

Missing	Missing	Censoring	Level of	ßc	Вмо	BMG	γc	Y _{Mo}	Y _{Mc}
outcome	confounder	0011001110	missingness	μ	PMO	РМС			
Yes	No	No	Low	3.113	3.215	3.720	0	0	1
No	Yes	No	Low	3.113	3.215	3.720	0	1	0
Yes	Yes	No	Low	3.113	3.215	3.720	0	0	0
Yes	No	Yes	Low	3.113	3.215	3.720	1	0	1
No	Yes	Yes	Low	3.113	3.215	3.720	1	1	0
Yes	Yes	Yes	Low	3.113	3.215	3.720	1	0	0
Yes	No	No	High	2.395	2.298	2.767	0	0	1
No	Yes	No	High	2.395	2.298	2.767	0	1	0
Yes	Yes	No	High	2.395	2.298	2.767	0	0	0
Yes	No	Yes	High	2.395	2.298	2.767	1	0	1
No	Yes	Yes	High	2.395	2.298	2.767	1	1	0
Yes	Yes	Yes	High	2.395	2.298	2.767	1	0	0



Low missing data (true ATE is 0.2)

Y	Х	Censoring	MI				IPW				
			Bias	MSE	Emp. SE	MC err.	Bias	MSE	Emp. SE	MC err.	
Yes	No	No	1.28%	0.0105	0.0102	0.0003	0.78%	0.0161	0.0160	0.0005	
No	Yes	No	1.20%	0.0099	0.0096	0.0003	1.58%	0.0194	0.0192	0.0006	
Yes	Yes	No	1.22%	0.0101	0.0098	0.0003	1.46%	0.0164	0.0161	0.0005	
Yes	No	Yes	2.12%	0.0115	0.0107	0.0003	1.38%	0.0137	0.0135	0.0004	
No	Yes	Yes	0.98%	0.0106	0.0104	0.0003	1.84%	0.0148	0.0143	0.0004	
Yes	Yes	Yes	1.46%	0.0110	0.0106	0.0003	1.68%	0.0139	0.0135	0.0004	

- MI has the lowest empirical SE, mean square error and Monte Carlo error of bias.



High missing data (true ATE is 0.2)

Y	х	Censoring	MI				IPW				
			Bias	MSE	Emp. SE	MC err.	Bias	MSE	Emp. SE	MC err.	
Yes	No	No	1.10%	0.0116	0.0114	0.0004	1.03%	0.0297	0.0297	0.0009	
No	Yes	No	1.17%	0.0099	0.0097	0.0003	2.33%	0.0477	0.0475	0.0015	
Yes	Yes	No	1.19%	0.0108	0.0105	0.0003	1.19%	0.0393	0.0392	0.0012	
Yes	No	Yes	3.01%	0.0149	0.0137	0.0004	1.44%	0.0241	0.0240	0.0008	
No	Yes	Yes	0.63%	0.0130	0.0129	0.0004	2.22%	0.0311	0.0308	0.0010	
Yes	Yes	Yes	1.42%	0.0135	0.0132	0.0004	2.31%	0.0250	0.0246	0.0008	

- MI has the lowest empirical SE, mean square error and Monte Carlo error of bias.
- Bias is consistently low for both methods under all scenarios.



- IPW straightforward to address censoring, and hence why commonly used in epidemiological/biostatistical studies
- Our simulations suggest that **even in simple settings** with intermittent missing outcome IPW provides **less precise** estimates
- Impact more pronounced when both outcome and confounder are missing and % missingness is relatively high
- In some settings IPW will be challenging, say more than one outcome and several confounders missing
- **MI is computationally expensive**, particularly if combined with non-parametric Bootstrap



What further simulated scenarios are warranted?

1. Allow more 'complex' MAR scenarios

- Allow Mc_t to be a function of Y, and
- Mo_t to be a function of X

2. Add 'Method 3' - use MI for both monotone and non-monotone missing.

- Don't expect much gain for MI
- Worth considering for completeness?