GLOBAL FOOTPRINTS OF MONETARY POLICIES

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Very Preliminary February 7, 2020

Abstract

We study the international transmission of the monetary policy of the two world's giants: China and the US. From East to West, the channels of global transmission differ markedly. US monetary policy shocks affect the global economy primarily through their effects on integrated financial markets, global asset prices, and capital flows. EMEs in particular see both a reduction in inflows and a surge in outflows when the market tide turns as a result of a US monetary contraction. Conversely, international trade, commodity prices and global value chains are the main channels through which Chinese monetary policy transmits worldwide. AEs with a strong manufacturing sector are particularly sensitive to these disturbances.

Keywords: Monetary Policy; Global Financial Cycle; International spillovers; US; China

JEL Classification: E44, E52, F33, F42

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We are grateful to Michael Howell at CrossBorder Capital Ltd and to the CPB Netherlands Bureau for Economic Policy Analysis for sharing their data with us. Rey thanks the ERC for financial support (ERC grant 695722). The views expressed in this paper are those of the authors and do not represent those of the Bank of England or any of its Committees.

Summary

Introduction The large and increasing interconnectedness of global real and financial markets, the emergence of Global Value Chains (GVC), and of a Global Financial Cycle (GFC, Rey, 2013), all provide fertile ground for international spillovers. In fact, the unprecedented intricacy of global networks generates potentially new dimensions for the international transmission of monetary policy shocks that go beyond the standard textbook trade channels primarily brought about by fluctuations in the exchange rate.

In a standard open economy environment a monetary tightening influences the current account in two ways. On the one hand, imports are affected by the contraction in domestic demand; on the other, the subsequent appreciation of the domestic currency makes exports more expensive, if one operates under a local currency pricing paradigm.¹ Foreign economies are affected by the local monetary policy shock only to the extent that their trade in goods and services depends on the local trade balance.

Miranda-Agrippino and Rey (2015) document a further channel for the international transmission of monetary policy that instead works through global financial markets. The synchronization of international financial markets epitomized by the GFC, together with the role of the dollar as the dominant currency of the international monetary system, consign a special role to US monetary policy as one of the drivers of the GFC itself (see also Jorda, Schularick, Taylor and Ward, 2018; Habib and Venditti, 2019). US monetary contractions are followed by a significant deleveraging of global financial intermediaries, a rise in aggregate risk aversion, a contraction in global asset prices and in global credit, a widening of corporate bond spreads, and a retrenchment in gross capital flows. The effects, estimated on the period 1980-2010 (or 1990-2010), are economically significant, and not confined only to countries that adopt an exchange rate peg.² Dées and Galesi (2019) confirm this result using a GVAR where trade-weights summarize the network

¹Important distortions here arise when one introduces dominant currency pricing that arises from USD invoicing as documented in Gopinath, Boz, Casas, Díez, Gourinchas and Plagborg-Møller (2020).

²Miranda-Agrippino and Rey (2020) extend the analysis to the unconventional monetary policy that the Fed adopted starting in 2009. US monetary policy that operates mainly on the short end of the yield curve elicits effects on global financial variables that are very similar to those documented for the pre-ZLB sample. One notable exception is the VIX, whose role as barometer of global risk may have changed since the crisis (see also Avdjiev, Gambacorta, Goldberg and Schiaffi, 2017; Forbes and Warnock, 2019; Burcu, Lombardi, Mihaljek and Shin, 2020).

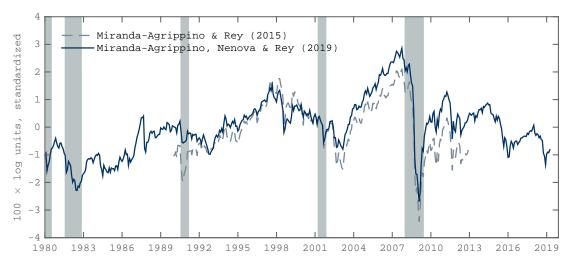
of cross-country interactions; US monetary policy shocks transmit across border almost irrespective of the exchange rate regime of the recipient country, and taking into account higher-order spillovers within the global trade network amplifies the transmission (see also Georgiadis, 2017).

But is the Federal Reserve the only giant capable of influencing global conditions through its policies?

In this paper we compare the global effects of US monetary policy with those elicited by surprise changes in the Chinese monetary policy stance. For the first time we are able to make use of a monetary policy index that summarizes the policy stance of the People's Bank of China, and can be regarded as the analog of the Federal Funds Rate (Xu and Jia, 2019). Postulating a standard Taylor-type rule for the Chinese monetary authority, and a delayed response for most of the global aggregates, we evaluate empirically how the world adjusts to a Chinese monetary policy shock.

Our estimates suggest that the monetary policies of the US and China have a sizeable impact on the global economy. However, the channels of transmissions of these spillovers differ. US shocks propagate predominantly through financial markets: financial conditions, risk indices, asset prices, private liquidity, and international capital flows all respond very significantly. Moreover, while tighter US monetary policy leads to a contraction of capital flows both in and out of the US, pointing to a general weakening of global financial activity, EMEs also suffer additional capital flights that contribute to increase their vulnerability. Conversely, global financial variables do not appear to be the primary transmission channel when one focuses instead on Chinese monetary policy shocks. In this case, it is the contraction of domestic demand and prices that drags down global activity. Commodity prices contract with some delay, but very significantly. Global asset prices are essentially insensitive for several months, after which they contract presumably as a result of the fall in commodity prices, and in global growth. Similarly, financial conditions significantly tighten for major commodity exporters, while they are largely unaffected at the global level. Commodity producers also experience capital flights and disinvestment. Global trade, and global growth as a consequence, contract. Through global value chains, these repercussions ripple all the way into Europe: German output, imports and exports all suffer severe contractions.





Note: Global factors in risky asset prices. Dashed line: factor in Miranda-Agrippino and Rey (2015). Solid line: update.

We summarize fluctuations in global asset prices and capital flows through global factors. For asset prices, we extend the global factor of Miranda-Agrippino and Rey (2015) along two dimensions: time, with estimates now covering all the months between 1980:01-2019:04, and cross-section, by performing the extraction on a larger and richer set of price series that is updated to reflect compositional changes in global markets, particularly through the inclusion of Chinese stocks. We then conduct a thorough factor analysis of global capital flows. Here too we find evidence of a dominant common global component that, very interestingly, but perhaps not surprisingly, strongly correlates with the global factor in asset prices, providing further additional evidence of the potency of the GFC.

Global Factors for Asset Prices and Capital Flows Miranda-Agrippino and Rey (2015) document the presence of a unique common factor in global risky asset prices that captures a significant share of common variation in global markets. In a simple model of heterogeneous financial intermediaries, this factor is primarily a function of realized market variance, and of the aggregate degree of time-varying risk aversion in global markets. The factor indeed displayed a significant correlation with other independent risk indices such as the VIX and other measures of implied variance.

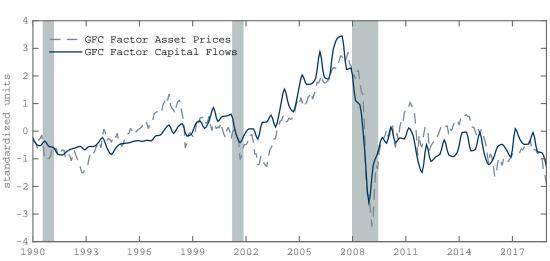


FIGURE 2: GLOBAL PRICES & CAPITAL FLOWS

Note: Dashed line: global factor in world risky asset prices. Solid line: global factor in world capital flows.

The factor was originally extracted from a set of monthly asset prices, all expressed in USD, up to 2010. Since then, the composition of global markets has changed, with Eastern markets gaining increasingly more visibility. In order to account for this, here we extend the analysis in two ways: (i) time, the new monthly factor covers the period from 1980:01 to 2019:04; (ii) cross-section, taking again as reference the components of the S&P Global index (https://us.spindices.com/indices/equity/sp-global-1200), we extract the factor from an updated sample that reflects this compositional changes, and includes more Chinese stocks. The methodology is the same as in Miranda-Agrippino and Rey (2015).

The old and new factors are plotted in Figure 1. Over the overlapping sample, the two factors are essentially the same. Since 2010, the factor picks up other important global events such as the struggle in European markets during the sovereign crisis; the global equity sell-off of the beginning of 2016, triggered by fears that the Chinese growth slowdown may have spiralled out of control, and by the dramatic plunge in oil prices; and the slowdown at the end of 2018, which the commentators attribute to the combined effect of the withdrawal of some monetary stimuli, and of the escalation in the US-China trade conflict.

Figure 2 compares the global factor in asset prices with a global factor in international capital flows. The correlation between the two is remarkable, and over 0.8. This serves

as additional evidence that international financial markets largely dance to the same tune. Our analysis complements that in Valente, Wincoop and Davis (2019) who also document important similarities between our asset prices factor and common factors in capital flows. Strikingly, they show that net flows as well as gross flows share a significant degree of common variation, that and the global factors they identify explain from 40% to half of their variance.^{3,4}

The capital flow factor is extracted from a large cross-section of data that we take from the IMF statistics, and cover all types of both in- and outflows (foreign direct investments, portfolio equity and debt, banking) across all countries.⁵ The data are quarterly at source, and we convert them to monthly via interpolation.⁶ The interpolation is useful for the purpose of constructing a monthly capital flows factor that we can compare with that in asset prices. It is however inconsequential for what concerns the study of the response of capital flows to monetary policy shocks: estimating responses using interpolated data or with alternative data that are available at monthly frequency (distributed by CrossBorder Capital Ltd.) deliver essentially the same results. We report additional details on the factors in Appendix B.

Global Transmission of US Monetary Policy Shocks We set the stage by looking again at the global transmission of US monetary policy shocks. This serves us as a benchmark to then discuss the Chinese responses that follow.

We summarize the global landscape with the following variables: world production and world trade (from the CPB World Trade Monitor); world financial conditions and world private liquidity (from CrossBorder Capital Ltd.); the global factors in asset prices and capital flows; exchange rates and a commodity price index. Data definitions and sources are collected in Appendix A.

We report median responses and posterior credible sets at the 68% and 90% level to a

³The first global factor in Valente et al. (2019) comoves closely with the global factor in asset prices of Miranda-Agrippino and Rey (2015), while the second is linked to commodity (energy) prices.

 $^{^{4}}$ Barrot and Serven (2018) reach similar conclusions although they document some degree of heterogeneity between AEs and EMEs. Conversely, Cerutti, Claessens and Rose (2017) report more conservative estimates.

⁵The estimate of the common factor is robust to using inflows and outflows separately, and to different specifications of the factor model.

⁶We interpolate level data using a shape-preserving piecewise cubic interpolation. Matlab command: y1 = interp1(t0,y0,t1,'pchip');.

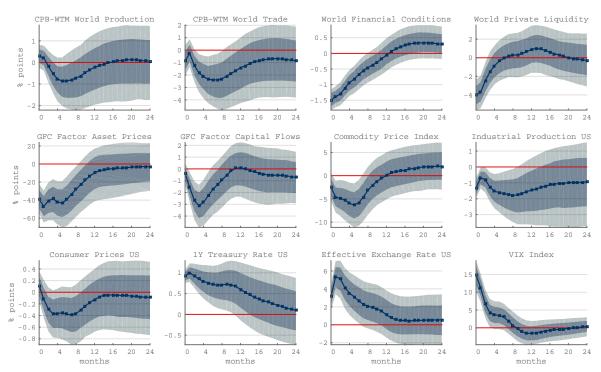


FIGURE 3: RESPONSES TO US MP SHOCK: #1

Note: Median IRFs with 68% and 90% posterior credible sets. BVAR(12). 1991:01-2018:12.

monetary policy shock identified using high-frequency movements in the price of Federal Funds Futures around FOMC announcements as external instrument (Stock and Watson, 2018). The IRFs are normalized such that the impact response of the policy rate (1-year rate) is equal to 100 bps. The VAR is estimated at monthly frequency over the sample 1991:01-2018:12 with 12 lags and standard macroeconomic priors (Giannone, Lenza and Primiceri, 2015).

Results are in Figures 3 and 4. Figure 3 largely replicates findings in Miranda-Agrippino and Rey (2015); following a US monetary policy tightening global financial conditions deteriorate materially. Private liquidity, measured as net credit generated by all credit providers, contracts. Global asset prices and global capital flows, summarized by the two factors, contract on impact, and the VIX spikes up. Global growth does not seem to be materially affected, while world trade contracts slightly at medium horizons. All this against a backdrop of cooling domestic conditions, with prices and production sliding. The US dollar appreciates (see also Degasperi, Hong and Ricco, 2019).

In Figure 4 we further explore the response of capital flows. We replace the global

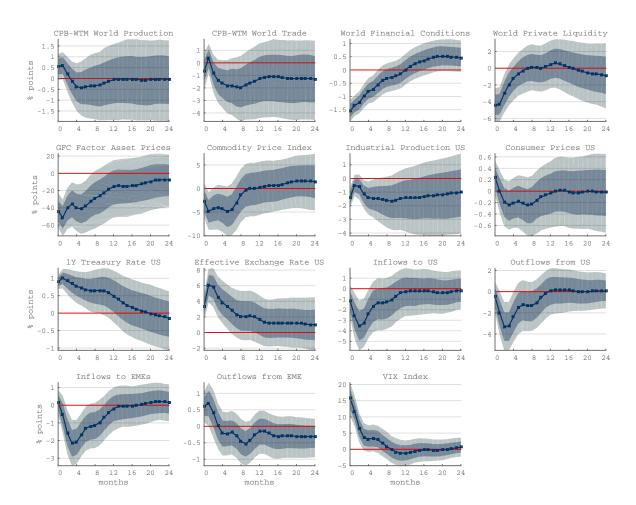


Figure 4: Responses to US MP Shock: #2

Note: Median IRFs with 68% and 90% posterior credible sets. BVAR(12). 1991:01-2018:12.

factor with four variables: capital flows in and out of the US; and capital flows in and out of EMEs. We note here that US inflows and outflows move largely in tandem, pointing to a general weakening of financial activity. This is not the case for EMEs, which are hit by a double whammy of less inflows and capital flights. This added vulnerability of EMEs to US monetary policy has been noted on several occasions, and culminated with the Taper Tantrum episode of 2013, when hints that the monetary stimulus may eventually be withdrawn threw investors into a mild panic that quickly transformed into excess volatility and sell-off, particularly in EMEs.

Global Transmission of Chinese Monetary Policy Shocks Particularly in recent years, the analysis of the conduct of monetary policy of the People's Bank of China has

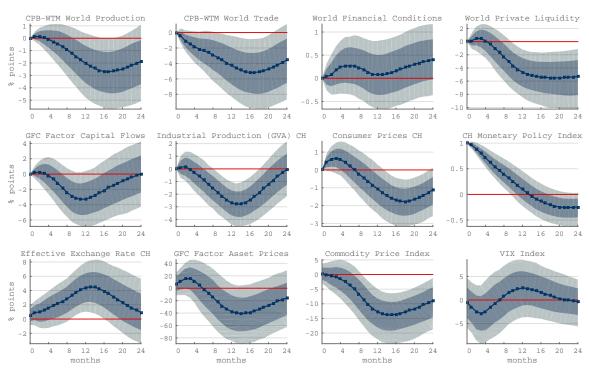


FIGURE 5: RESPONSES TO CHINESE MP SHOCK: #1

Note: Median IRFs with 68% and 90% posterior credible sets. BVAR(12). 1999:01-2018:12.

gained increasing attention (see e.g. Jones and Bowman, 2019). The objective of the prudent monetary policy of the Chinese monetary authority, initiated in 1989, is that of maintaining prices and the value of the Renminbi stable, while contributing to and promoting economic growth (Zhou, 2015). Over the years, the policy has moved from being predominantly quantity-based to interest-rate-based (Chen, Chen and Gerlach, 2011; Kim and Chen, 2019). And much like for other major central banks, communication has become increasingly important and studied (McMahon, Schipke and Li, 2018).

To measure the Chinese monetary policy stance, we make use of the monetary policy indicator constructed in Xu and Jia (2019) to summarize information in a variety of interest rates. We identify Chinese monetary policy shocks by postulating a Taylor-type rule for the monetary authority, as an innovation of the monetary policy index in a recursively identified VAR. Together with domestic prices and output, we assume that world variables do not react within a month. The VARs are monthly, estimated with 12 lags from 1999:01 to 2018:12, and IRFs are normalized to yield a 1% increase in the monetary policy index on impact. The sample standard deviation of the index is 0.5. The

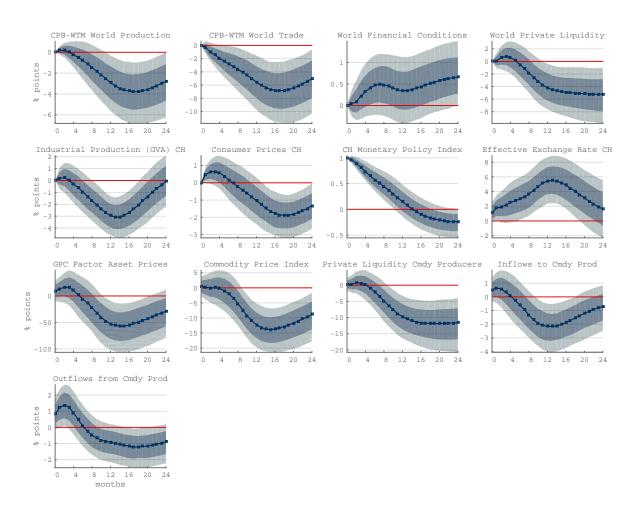


FIGURE 6: RESPONSES TO CHINESE MP SHOCK: #2

Note: Median IRFs with 68% and 90% posterior credible sets. BVAR(12). 1999:01-2018:12.

normalization can thus be thought of as a two standard deviations shock, hence quite large.

Figure 5 evaluates the global effects of Chinese monetary policy shocks against the same set of global variables of Figure 3. Following the shock, the monetary policy indicator monotonically returns to trend after about 15 months. Chinese production, measured as gross value added, declines with delay, reaching a peak negative response after one year. Similar dynamics characterize the price adjustment and the reaction of the RMB. Prices eventually decline, while the currency slowly appreciates. The domestic response is very much in line with the standard textbook transmission mechanism documented for other countries, apart from the slow exchange rate adjustment. The channels of global

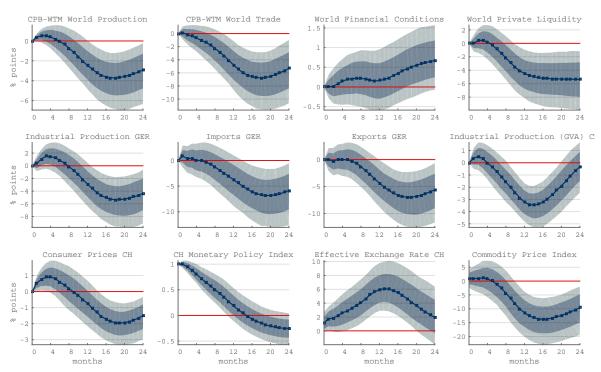


FIGURE 7: RESPONSES TO CHINESE MP SHOCK: #3

Note: Median IRFs with 68% and 90% posterior credible sets. BVAR(12). 1999:01-2018:12.

transmission are instead very different from those documented for the US. Global financial variables are largely unaffected: world financial conditions, the VIX, and the global factors in asset prices and capital flows do not respond in any significant way, at least at short-medium horizons. Conversely, world production slows down, presumably dragged by the contraction in Chinese domestic demand that in turn pulls down world trade and commodity prices. World private liquidity eventually declines, potentially a result of the slowdown in global growth. The sluggish response of the exchange may potentially hold the key for the difference in the global spillovers of the Chinese monetary policy shocks relative to the US ones, and deserves further investigation (see also Richmond, 2019).

The effects that fluctuations in the Chinese economy elicit on global quantities seems to go mainly through commodity prices, and the compression of global demand. We explore this further in Figures 6 and 7. In Figure 6 we look more in detail at how commodity producers react to the shock. Financial conditions tighten significantly for this pool of countries; in particular, they witness both a contraction in inflows and a surge of capital outflows. Hence, an important channel of the international transmission of Chinese monetary shocks seems to reside in its large relative weight in world production. Weak Chinese demand has the potential to disrupt global production because of the crucial role it plays in the global markets for both raw materials, and intermediate production goods. As a consequence, AEs whose economy is particularly reliant on manufacturing production, and that operate across multiple GVCs, may be particularly sensitive to these types of disturbances. In Figure 7 we look in particular at the case of Germany. Our results show that German imports and exports both contract significantly, with consequential detrimental effects on production.

Conclusions We compare the global transmission of the monetary policy of the two world's giants: the US and China. We find that both have a significant global footprint, but that they operate through fundamentally different channels.

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| CODE | Variable | Source |
|------------|--|-------------------------------------|
| WORLDIP | World Production, Merges Old and New Data, Base Year 2010=100, Prod. Weights | CPB-Wolrd Trade Monitor + Own Calc. |
| WORLDTRADE | World Total Merchandise Trade, Volume Index, Merges Old and New Data | CPB-Wolrd Trade Monitor + Own Calc. |
| WORLDFCI | World Financial Conditions, Diffusion Index | CrossBorder Capital |
| WORLDPLIQ | World Private Liquidity, Stock, USD | CrossBorder Capital |
| INDPRO | US Industrial Production Index, Index 2012=100, SA | St Louis Fed FRED Database |
| CPIAUCSL | US Consumer Price Index for All Urban Consumers: All Items, Index 1982-1984=100, SA | St Louis Fed FRED Database |
| GS1 | US 1-Year Treasury Constant Maturity Rate, Percent, NSA | St Louis Fed FRED Database |
| GERINDPRO | Production of Total Industry in Germany, Index 2015=100, Monthly, SA | St Louis Fed FRED Database |
| GERIMP | Imports: Value Goods for Germany, National currency Billions, Monthly Level, SA | St Louis Fed FRED Database |
| GEREXP | Exports: Value Goods for Germany, National currency Billions, Monthly Level, SA | St Louis Fed FRED Database |
| CHINDPRO | China Real Gross Value Added (Industrial Production), NSA, year-on-year % change | China National Bureau of Statistics |
| CHCPI | China Consumer Price Index, Not SA, year-on-year percentage change | China National Bureau of Statistics |
| CHMPI | Chinese Monetary Policy Index | Xu and Jia (2019) |
| USDEER | United States Real Effective ExhangeRate, Narrow Index | Bank for International Settlements |
| RMBEER | China Real Effective ExhangeRate, Broad Index | Bank for International Settlements |
| CRBPI | Commodity Research Bureau Price Index, Spot Prices, End of Month | Datastream |
| GFCFAC | Common Factor in Risky Asset Prices, 2019 Vintage, Standardized | Own Calculations |
| VIX | CBOE Volatility Index, Annualized Percentage Points, NSA | Datastream |
| CAPFGF | Global Factor Capital Flows, All Flows Types, Inflows and Outflows, Interp. | IFS, BOPS, WEO + Own Calculations |
| OUTFEME | Outflows from Emerging Markets, Percentage of GDP, All Flows Types, Interp. | IFS, BOPS, WEO + Own Calculations |
| INFLEME | Inflows to Emerging Markets, Percentage of GDP, All Flows Types, Interpolated | IFS, BOPS, WEO + Own Calculations |
| OUTFCMDY | Outflows from Major Commodity Producers, Percentage of GDP, All Flows Types, Interp. | IFS, BOPS, WEO + Own Calculations |
| INFLCMDY | Inflows to Major Commodity Producers, Percentage of GDP, All Flows Types, Interpolated | IFS, BOPS, WEO + Own Calculations |
| OUTFUS | Outflows from US, Percentage of GDP, All Flows Types, Interpolated | IFS, BOPS, WEO + Own Calculations |
| INFLUS | | IFS, BOPS, WEO + Own Calculations |
| PLIQCMDY | Private Liquidity Major Commodity Producers, Stock, USD | CrossBorder Capital |

TABLE A.1: Data Series in Bayesian VARs

A Data Appendix

B Appendix on Global Financial Factors

We estimate global factors in (i) asset prices and (ii) international capital flows using a dynamic factor model first applied to asset prices in Miranda-Agrippino and Rey (2015). Tables B.1 and B.2 describe the underlying series that are used in the estimation of these factors.

The first factors extracted respectively from all asset prices and all capital flows are plotted in Figure 2 in the main text. There is a high correlation between the asset price and capital flow factors despite the fact that the underlying datasets used in their construction do not overlap.

Table B.4 reports the share of the variance in underlying asset price and capital flow series that the first two factors explain. For the capital flows factor, we also report the relevant statistics for the second and third factors (the equivalent statistics) as well as the correlation between the global asset price and capital flow factors.

| Asset Class | Index Universe | DETAILS |
|-------------|-----------------|--|
| Equity | S&P Global 1200 | Full set of total return in- |
| | | dex constituents as of May |
| | | 2019. For index construction |
| | | <pre>see https://us.spindices.com/</pre> |
| | | indices/equity/sp-global-1200. |
| Commodities | Datastream | Prices of 126 different commodities in- |
| | | cluding oil, gas, agricultural commodi- |
| | | ties, mining, non-precious metals. |
| Bonds | iBoxx | Total returns on 80 corporate bond in- |
| | | dices for Euro and Sterling markets. |
| | FTSE | Total returns on 150 corporate bond |
| | | indices from the WorldBIG, EuroBIG, |
| | | USBIG index series. |

TABLE B.1: Asset price data for global factor in asset prices

Capital flow aggregates used in the main analysis are constructed as follows:

Advanced economies: Indicated by "AE" in final column of Table B.2. We rely on the IMF classification of advanced economies as of end-2018.

Emerging economies: Indicated by "EME" in final column of Table B.2. We rely on the IMF classification of emerging and developing economies as of end-2018.

Commodity producers: Australia, Brazil, Canada, Chile, Colombia, Indonesia, Kazakhstan, Malaysia, Mongolia, New Zealand, Norway, Russia, Saudi Arabia, South Africa, Venezuela. Country grouping is based on that used by CrossBorder Capital for their regional liquidity and financial conditions indices.

Oil producers: Azerbaijan, Colombia, Kazakhstan, Norway, Russia, Saudi Arabia, Venezuela. Classified according to the average net exports of fuel over 1995–2018 from UNCTAD trade statistics by product groups.

| COUNTRY | FDI | PE | PD | OTH | AE or EME? |
|----------------|--------------|--------------|--------------|--------------|------------|
| Albania | \checkmark | | | \checkmark | EME |
| Argentina | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Armenia | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Australia | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Austria | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Azerbaijan | \checkmark | | | \checkmark | EME |
| Bangladesh | \checkmark | | | \checkmark | EME |
| Belarus | \checkmark | | | \checkmark | EME |
| Belgium | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Bosnia and | \checkmark | | | \checkmark | EME |
| Herzegovina | | | | | |
| Brazil | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Bulgaria | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| CANADA | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| CHILE | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| China | \checkmark | | | \checkmark | EME |
| Colombia | \checkmark | | | \checkmark | EME |
| Costa Rica | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Croatia | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Cyprus | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Czech Republic | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Denmark | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Ecuador | \checkmark | | | \checkmark | EME |
| El Salvador | \checkmark | | | \checkmark | EME |
| Estonia | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Finland | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| France | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Georgia | \checkmark | | | \checkmark | EME |
| Germany | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Greece | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Guatemala | \checkmark | | | \checkmark | EME |
| Hong Kong | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| HUNGARY | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Iceland | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| India | \checkmark | | | \checkmark | EME |
| Indonesia | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Ireland | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| ISRAEL | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Italy | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Japan | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Jordan | \checkmark | | | \checkmark | EME |
| Kazakhstan | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Korea | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| | | | | | |

TABLE B.2: Country and instrument list for global factor in capital flows

| COUNTRY | FDI | PE | PD | OTH | AE or EME? |
|----------------|--------------|--------------|--------------|--------------|------------|
| LATVIA | | \checkmark | \checkmark | <u>√</u> | AE |
| LEBANON | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| LITHUANIA | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| LUXEMBOURG | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Malaysia | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Malta | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Mauritius | \checkmark | | | \checkmark | EME |
| Mexico | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Mongolia | \checkmark | | | \checkmark | EME |
| Montenegro | \checkmark | | | \checkmark | EME |
| Morocco | \checkmark | | | \checkmark | EME |
| Namibia | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Netherlands | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| New Zealand | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| N. Macedonia | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Norway | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Pakistan | \checkmark | | | \checkmark | EME |
| Panama | \checkmark | | | \checkmark | EME |
| Peru | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Philippines | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Poland | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Portugal | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Russia | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Saudi Arabia | \checkmark | | | \checkmark | EME |
| Serbia | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Singapore | \checkmark | | | \checkmark | AE |
| Slovak Rep. | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Slovenia | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| South Africa | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| Spain | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| Sri Lanka | \checkmark | | | \checkmark | EME |
| Sweden | \checkmark | | | \checkmark | AE |
| Switzerland | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| THAILAND | \checkmark | \checkmark | \checkmark | \checkmark | EME |
| TURKEY | \checkmark | | | \checkmark | EME |
| UNITED KINGDOM | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| UNITED STATES | \checkmark | \checkmark | \checkmark | \checkmark | AE |
| URUGUAY | \checkmark | | | \checkmark | EME |
| VENEZUELA | \checkmark | \checkmark | \checkmark | \checkmark | EME |

 TABLE B.3: Country and instrument list for global factor in capital flows

 (CONTINUED)

Note: FDI stands for foreign direct investment flows, PE stands for portfolio equity investment flows, PD stands for portfolio debt investment flows, and OTH stands for other investment flows. AE is short for Advanced Economy; EME is short for Emerging Market Economy. We include both in- and outflows for the respective countries and instrument.

| Factor | % Covariance | % Spectral | IC_p1 | IC_p2 | IC_p3 | ONATSKI |
|--------|--------------|------------|---------|---------|---------|---------|
| | MATRIX | DENSITY | | | | Test |
| GFCFAC | 21.5% | 24.1% | -0.184 | -0.183 | -0.189 | 0.049 |
| CFFGU1 | 7.8% | 20.7% | -0.042 | -0.040 | -0.049 | 0.041 |
| CFFGU2 | 5.1% | 14.5% | -0.051 | -0.047 | -0.065 | 0.007 |
| CFFGU3 | 4.4% | 12.0% | -0.055 | -0.049 | -0.076 | 0.988 |

TABLE B.4: Global factors in asset prices and capital flows

Note: The first column of the table shows the % of variance explained by the eigenvalue corresponding to each factor of the covariance matrix of the respective data (either asset prices or capital flows). The second reports the % of variance explained by the same eigenvalue of the spectral density matrix of the data. The following three columns report the value of the IC_p criteria in Bai and Ng (2002) and the last shows the p-value for the Onatski (2009) test where the null of r - 1 common factors is tested against the alternative of r common factors.

| | GFCFAC | CFFGU1 | CFFGU2 | CFFGU3 |
|------------|--------|--------|--------|--------|
| GFCFAC | 1.00 | | | |
| CFFGU1 | 0.81 | 1.00 | | |
| CFFGU2 | 0.19 | 0.02 | 1.00 | |
| CFFGU3 | -0.06 | -0.04 | 0.03 | 1.00 |
| VIX | -0.28 | -0.20 | -0.24 | -0.48 |
| USDEER | -0.30 | -0.04 | -0.15 | 0.38 |
| GS2 | 0.17 | 0.21 | -0.62 | -0.07 |
| GS10 | 0.05 | 0.09 | -0.66 | -0.15 |
| WORLDFCI | -0.50 | -0.43 | -0.21 | 0.17 |
| WORLDPLIQ | 0.09 | 0.00 | 0.83 | 0.08 |
| CRBPI | 0.19 | -0.08 | 0.92 | 0.04 |
| OILP | 0.28 | 0.15 | 0.82 | -0.08 |
| WORLDIP | 0.09 | 0.05 | 0.77 | 0.18 |
| WORLDTRADE | 0.16 | 0.14 | 0.73 | 0.12 |
| SHIPPING | 0.25 | 0.26 | 0.26 | 0.25 |

TABLE B.5: Correlations between global factors in asset prices and capital flows and key financial & economic variables

Note: For variable definitions see Table A.1 in Appendix A. The correlation coefficients above relate to the full sample period of the capital flow factors (1990-2018) or the longest possible sub-period for which the covariates are available.