Power laws in market microstructure

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- Characterisation of equilibrium
- Existence and further properties
- Impact asymptotics and trade volume

6 Numerics



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- Models used by practitioners include square root and logarithm (e.g. Torre (1997), Potters and Bouchaud (2003), Almgren et al. (2005), Bershova and Rakhlin (2013), and Zarinelli et al. (2015)...).

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More empirical evidence

Several studies (Gopikrishnan et al. (2000), Lillo et al. (2005), Vaglica et al. (2008), Bershova and Rakhlin (2013)...) showed that metaorders have tail distribution following a power law (with exponents ranging from 1.56 to 1.74!).

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- To introduce non-uniform pricing one should consider a limit order market as in Glosten (1994) (*Is the electronic open limit order book inevitable?*, J. of F.).
- However, Biais, Hillon and Spatt (1995) and Sandas (2001) show that the empirical findings strongly contradict the predictions of many microstructure models on limit order markets including Glosten (1994).

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- Trading takes place at t = 0 and t = 1.
- Market consists of a riskless asset with r = 0 and a single risky asset. The fundamental value of the asset *V* will be revealed to the public at time 1.

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- *N* Informed investors know *V* and are risk-neutral, i.e. they maximise their expected wealth at time 1.
- A trading desk receiving orders from noise and informed traders. The desk does not trade in its own account and thus a market order of size *y* is priced at

$$\int_{Y}^{Y+y} h(x) dx,$$

where *Y* is the accumulated number of shares from earlier trades.

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 Since h is nondecreasing, the first order condition characterises the optimal X* via V = F(X*), where

$$F(x) := \int_{-\infty}^{\infty} h(x+z)q(\sigma,z)dz$$
 (1)

and $q(\sigma, \cdot)$ is the probability density function of a mean-zero Gaussian random variable with variance σ_{ϵ}^2 .

The case N > 1

- Assume all informed orders arrive after the noise.
- As every insider has symmetric information and is risk-neutral, in a symmetric equilibrium, the demand x* for each insider must be the same and satisfy

$$v=E^{\nu}\left[\frac{h(Z+Nx^*)}{N}+\frac{N-1}{N^2x^*}\int_0^{Nx^*}h(Z+u)du\right].$$

• Denoting the total informed demand by X^* , the above can be rewritten as $V = F(X^*)$, where

$$F(x) := E^{\nu} \left[\frac{h(Z+x)}{N} + \frac{N-1}{Nx} \int_0^x h(Z+u) du \right], \quad (2)$$

and F(0) is interpreted by continuity to be

$$E^{\nu}\left[\frac{h(Z)}{N}+\frac{(N-1)h(Z)}{N}\right]=E^{\nu}[h(Z)].$$

The limit order book and equilibrium

Following Glosten, we assume limit prices are given by 'tail expectations:

$$h(y) = \left\{ \begin{array}{l} E[V|Y \ge y], & \text{if } y > 0; \\ E[V|Y \le y], & \text{if } y < 0. \end{array} \right\}$$
(3)

Definition 1

The pair (h^*, X^*) is said to be a Glosten equilibrium if h^* is non-decreasing and non-constant, $X^* \in \mathbb{R}$ and

- i) h^* satisfies (3) with $Y = X^* + Z$;
- ii) X^* is the profit maximising order size for the insider(s) given h^* . That is, $V = F(X^*)$, where F is given by (2).

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Few objects of interest

- Suppose that (X*, h*) is an equilibrium and write h instead of h* to ease exposition.
- Introduce the functions Φ^{\pm} and Π^{\pm} via

$$\begin{array}{rcl} \Phi^+(y) & := & E[V\mathbf{1}_{[V>y]}], \ \Pi^+(y) := P(V>y) \\ \Phi^-(y) & := & E[V\mathbf{1}_{[V\leq y]}], \ \Pi^-(y) := P(V\leq y) = 1 - \Pi^+(y). \end{array}$$

- Define $\Psi^{\pm}(y) := \frac{\Phi^{\pm}(y)}{\Pi^{\pm}(y)}$ so that $\Psi^{+}(y) = E[V|V > y]$ and $\Psi^{-}(y) = E[V|V \le y].$
- Since, for y > 0, $h(y) = E[V|F^{-1}(V) + Z \ge y]$,

$$h(y) = E[V|V \ge F(y-Z)] = \frac{E[V\mathbf{1}_{[V \ge F(y-Z)]}]}{P(V \ge F(y-Z))}$$
$$= \frac{\int_{-\infty}^{\infty} \Phi^{+}(F(y-z))q(\sigma,z)dz}{\int_{-\infty}^{\infty} \Pi^{+}(F(y-z))q(\sigma,z)dz} (\neq E[\Psi^{+}(F(y-Z))]!)$$

An analogous representation holds for y < 0.

An equation for F

• Define, for any continuous *g*, the mappings

$$\phi_g^{\pm}(x) := \frac{\int_{-\infty}^{\infty} \Phi^{\pm}(g(z))q(\sigma, x - z)dz}{\int_{-\infty}^{\infty} \Pi^{\pm}(g(z))q(\sigma, x - z)dz}$$

Let us also set

$$\phi_g(x) := \phi_g^+(x) \mathbf{1}_{x \ge 0} + \phi_g^-(x) \mathbf{1}_{x < 0}.$$
(4)

Now, combining all of the above yields an equation for F:

$$F(x) = \frac{1}{N} \int_{-\infty}^{\infty} q(\sigma, x - z) \phi_F(z) dz \qquad (5)$$

+ $\frac{N - 1}{Nx} \int_0^x dy \int_{-\infty}^{\infty} q(\sigma, y - z) \phi_F(z) dz.$

Given the above consideration the following now is obvious:

Theorem 2

Equilibrium exists if and only if there exists a function $F : \mathbb{R} \to \mathbb{R}$ that satisfies (5). Given such a solution F, (X^*, h^*) constitutes an equilibrium, where $X^* = F^{-1}(V)$ and h^* is defined via (4) and its counterpart for y < 0.

Therefore, finding an equilibrium boils down to finding a solution of (5).

Examples

• Suppose $P(V = 1) = P(V = -1) = \frac{1}{2}$. Then, the unique symmetric solution of (5) is defined for x > 0 by

$$F(x) = \frac{1}{N} \int_0^\infty q_0(\sigma, x, z) dz + \frac{N-1}{Nx} \int_0^x dy \int_0^\infty dz q_0(\sigma, y, z),$$

where $q_0(\sigma, y, z) := q(\sigma, y - z) - q(\sigma, y + z)$. Moreover, $X^* = \infty$ (resp. $X^* = -\infty$) if V = 1 (resp. V = -1) and, thus, $h^*(y) = \mathbf{1}_{[y>0]} - \mathbf{1}_{[y<0]}$. Nevertheless, insiders' profit remains finite:

$$\int_0^\infty E^1(1-h(Z+y))dy = 2E^1\left(\int_0^\infty \mathbf{1}_{[Z<-y]}dy\right) = \sigma\sqrt{\frac{2}{\pi}}.$$

Examples

• Suppose $P(V = -1) = P(V = 0) = P(V = 1) = \frac{1}{3}$. Then, similar considerations yield

$$F(x) = \frac{1}{N} \int_0^\infty q_0(\sigma, x, z) \frac{1}{1 + P(Z \ge z)} dz$$
$$+ \frac{N - 1}{Nx} \int_0^x dy \int_0^\infty dz q_0(\sigma, y, z) \frac{1}{1 + P(Z \ge z)}.$$

Again, $X * (1) = \infty$ and $X(-1) = -\infty$. But $X^*(0) = 0$. Consequently, the order book will not be flat. In particular, for y > 0

$$h(y) = \frac{E[V\mathbf{1}_{[X^*(V)+Z \ge y]}]}{P(X^*(V)+Z \ge y)} = \frac{P(V=1)}{P(V=1)+P(V=0,Z \ge y)}$$
$$= \frac{1}{1+P(Z \ge y)}.$$

Moreover, the bid-ask spread is given by $h(0+) - h(0-) = \frac{4}{3}$, independent of the noise variance.

Scaling property and uniqueness

- Due to the scaling property of *q* one should expect *F* exhibit similar scaling properties.
- Indeed, if F(1; x) is a solution of (5) with σ = 1, then straightforward manipulations yield F(1; ^x/_σ) solves (5).
- Thus, if (5) has a unique solution for one *σ*, it has a unique solution for all.
- This scaling property is also inherited by *h*: if (5) has a unique solution, *h*(σ; *x*) = *h*(1; ^{*x*}/_σ) for all *x* ≠ 0. As a consequence, *X*^{*}(σ) = σ*X*^{*}(1).

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Consequences of uniqueness

- The spread, i.e. h(0+) h(0-), is independent of $\sigma!$
- The spread associated with trade size y > 0, i.e.
 h(y) h(-y), and, therefore, the aggregate mid-spread S is decreasing with the amount of noise trading, consistent with the experimental findings of Bloomfield et al. (2009).

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- Thus, the order book gets flatten as σ increases and converges to a model with proportional transaction costs.

Towards existence

 Let's denote the interior of the support of V by (m, M), where −∞ ≤ m < M ≤ ∞, and recall on the support of V

► Reminder
$$\Psi^{\pm}(y) = \frac{\Phi^{\pm}(y)}{\Pi^{\pm}(y)}$$
 (6)

so that $\Psi^+(y) = E[V|V > y]$ and $\Psi^-(y) = E[V|V \le y]$.

• For any continuous *g* let *u*⁺ (resp. *u*⁻) be the unique solution of

$$u_t + \sigma^2 u_{xx} = 0,$$
 $u(1, x) = \Pi^+(g(z)) \text{ (resp. } \Pi^-(g(z))\text{).}$ (7)

Let g be as above. Then the following hold:

There exits a solution B on a filtered probability space (Ω, F, (F_t), Q) to the following SDE:

$$dB_t = \sigma dW_t + \sigma^2 \frac{u_x(t, B_t)}{u(t, B_t)} dt, \quad B_0 = x,$$
(8)

where *u* is either u^+ or u^- and *W* is a Brownian motion with $W_0 = 0$.

φ⁺_g(x) = ℝ^{Q⁺} [Ψ⁺(g(B₁))] and φ⁻_g(x) = ℝ^{Q⁻} [Ψ⁻(g(B₁))], where (B, Q⁺) (resp. (B, Q⁻)) corresponds to the solution of (8) if u = u⁺ (resp u = u⁻) and ℝ^Q stands for the expectation under Q. ► Equation for *F*

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The following properties allow us to ensure in particular that the solution of (5) is increasing.

- $\phi_g^+(0) > \phi_g^-(0).$
- Suppose further that g is non-decreasing. Then, ϕ_g^{\pm} are non-decreasing, too. Consequently, ϕ_g is non-decreasing. Moreover,

$$\phi_g^+(x) \leq \mathbb{E}^{\mathbb{Q}^+}\left[\Psi^+(g(\sigma W_1 + x))\right]$$
(9)

$$\phi_{g}^{-}(x) \geq \mathbb{E}^{\mathbb{Q}^{-}}\left[\Psi^{-}(g(\sigma W_{1}+x))\right].$$
(10)

Theorem 3

Suppose $-\infty < m < M < \infty$. Then, there exists a Glosten equilibrium.

The above theorem is proved by means of Schauder's fixed point theorem, hence no claim of uniqueness is given.

Asymptotics for F and h

- Although it is not possible to find explicitly *F*, it is possible to obtain its asymptotics.
- Recall that g : (0,∞) → (0,∞) is said to be regularly varying of index ρ at ∞ if

$$\lim_{\lambda o\infty}rac{g(\lambda x)}{g(\lambda)}=x^
ho,\quad orall x>0.$$

Regular variation at $-\infty$ is defined analogously.

• It can be shown that *F* and *h* have the same regular variation index.

F is regularly varying

Suppose that N > 1, $-\infty < m < M < \infty$, and Π^+ has a continuous derivative. Note that $\Psi^+_x(M) := \frac{d\Psi^+(M-)}{dx} \le 1$.

• Then, M - F is regularly varying at ∞ with index

$$\rho^{+} = \frac{\Psi_{x}^{+}(M) - 1}{1 - \frac{\Psi_{x}^{+}(M)}{N}}.$$
(11)

• Under above assumptions F - m is regularly varying at $-\infty$ with index

$$\rho^{-} = \frac{\Psi_{x}^{-}(m) - 1}{1 - \frac{\Psi_{x}^{-}(m)}{N}}.$$
 (12)

The case of slow variation

- The above shows that if $\Psi_x^+(M) < 1, -1 < \rho^+ < 0$ and $M F(x) \sim x^{\rho^+}$ for large *x*.
- On the other hand, if Ψ⁺_x(M) = 1, F, hence h, is slowly varying at ∞.
- To obtain a better understanding of how slow the variation of *M* - *F* is, suppose that there exists an integer *n* and constant *k* > 0 such that

$$\frac{\Psi^+(x)-x}{(M-x)^n} \to \frac{1}{k} \quad \text{as } x \to M.$$
 (13)

• Then, it can be shown that

$$M-F(x)\sim \left(\frac{N}{N-1}\frac{k}{n}\right)^{\frac{1}{n}}(\ln x)^{-\frac{1}{n}}.$$

Distribution of the volume

• Note for *x* > 0

$$P(X^* > x) = P(F^{-1}(V) > x) = P(V > F(x)) = \Pi^+(F(x)).$$

 However, assuming (13) also yields Π⁺(F) is regularly varying. Thus,

$$P(X^* > x) = x^{-\zeta^+} s(x),$$

where s is a slowly varying function and

$$\zeta^+ := \frac{\Psi^+_x(M)}{1 - \frac{\Psi^+_x(M)}{N}}$$

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Distribution of the volume

Moreover, since Y* = X* + Z and Z and V are independent, we have for y > 0

$$P(Y^* > y) = \int_{-\infty}^{\infty} dz P(X^* > z) q(\sigma, y - z),$$

which is regularly varying at infinity with the same index.

Thus,

$$P(Y^* > y) = y^{-\zeta^+} s(y), \quad y > 0,$$
 (14)

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for some regularly varying *s*. In particular, if *V* has light tails, i.e. $\Psi_X^+(M) = 1$, $P(Y^* > y)$ is regularly varying of index $-\frac{N}{N-1}$. Fattalls

General signals

Although the above theory is currently limited to bounded V, formal calculations in the general case show that F is regularly varying at ∞ of order ρ⁺, where

$$\rho^{+} = \frac{\Psi_{x}^{+}(\infty) - 1}{1 - \frac{\Psi_{x}^{+}(\infty)}{N}}.$$
(15)

 However, since ρ⁺ must be non-negative, this places the restriction on N:

$$N > \Psi_X^+(\infty)$$
 (16)

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unless $\Psi_{X}^{+}(\infty) = 1$.

 That is, equilibrium requires a sufficiently large amount of competition among insiders when the asset value has fat tails. Table: Distributions with power-law impact

Distribution	Density	$ ho^+$	
Beta prime	$x^{\lambda-1}(1+x)^{-(\lambda+lpha)}$	$\left(\frac{N-1}{N}\alpha-1\right)^{-1}$	
Fréchet	$(x - \beta)^{-(1+\alpha)} \exp\left\{-\left(\frac{x-\beta}{s}\right)^{-\alpha}\right\}$	$\left(\frac{N-1}{N}\alpha-1\right)^{-1}$	
Lomax	$\left(1+\frac{x}{\lambda}\right)^{-(\alpha+1)}$	$\left(\frac{N-1}{N}\alpha-1\right)^{-1}$	
Pareto	$x^{-(lpha+1)}$	$\left(\frac{N-1}{N}\alpha-1\right)^{-1}$	
Student	$\left(1+\frac{x^2}{\alpha}\right)^{-(\alpha+1)/2}$	$\left(\frac{N-1}{N}\alpha-1\right)^{-1}$	
In above probability densities are given up to a scaling factor and implicit			

constraints are enforced to ensure they are well defined with finite mean. Moreover, $N > \frac{\alpha}{\alpha-1}$ in all of the above.

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Table: Distributions with logarithmic impact

Distribution	Density	Asymptotics		
Exponential	$\exp(-\lambda x)$	$\frac{N}{\lambda(N-1)}\log X$		
Gaussian	$\exp(-(x-\mu)^2/\Sigma)$	$\sqrt{\frac{2\Sigma N}{N-1}}\sqrt{\log X}$		
Inverse Gaussian	$x^{-3/2} \exp\left(-\frac{\lambda(x-\mu)^2}{2\mu^2 x}\right)$	$\frac{2N\mu^2}{\lambda(N-1)}\log X$		
Normal Inverse Gaussian	$\frac{K_1(\lambda\zeta(x))}{\pi\zeta(x)}\exp(\delta\gamma+\beta(x-\mu))$	$\frac{N}{(N-1)(\lambda+\beta-1)}\log X$		
Weibull	$x^{d-1} \exp(-\lambda^{p} x^{p})$	$\left(\frac{N}{\lambda^p(N-1)}\right)^{1/p} (\log x)^{1/p}$		
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In above probability densities are given up to a scaling factor and implicit constraints are enforced to ensure they are well defined with finite mean. Moreover, $\zeta(x) := \delta^2 + (x - \mu)^2$ for the Normal Inverse Gaussian distribution.

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Equilibrium with Student signals

Equilibrium: Student signals



Figure: Equilibrium solutions for Student signals for the cases N = 2, N = 3 and N = 25.

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Convergence to square root impact



Figure: Functional form of the equilibrium for Student signals for $\alpha = 3$, N = 25.

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Equilibrium with log-Normal signals



Equilibrium: Log-normal signals

Figure: The discontinuity of h(x) at the origin is the bid-ask spread.

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Conclusion

- Solved in a one-period setting the equilibrium in a limit order market in the framework of Glosten (1994).
- Although the equilibrium is not explicitly solvable, we show that the impact of large trades show regular variation and find the exact exponent.
- The impact seems to be of power law for asset values exhibiting fat tails and logarithmic for the ones with light tails. This provides a testable answer to the debate among the practitioners on the nature of the price impact.
- The equilibrium volume for the assets is of regular variation. Thus, the model can be seen as a justification in an REE setting to several conclusions of the econophysics literature.

 Suppose that Π⁺(x) has a power-like behaviour in the sense that

$$\frac{\Pi_x^+(x)}{\Pi^+(x)} \sim -\alpha (M-x)^{-1}, \quad \alpha > 0.$$

An application of L'Hospital rule shows that

$$\frac{\int_{x}^{M} \Pi^{+}(y) dy}{\Pi^{+}(x)(M-x)} \sim \frac{\Pi^{+}(x)}{\Pi^{+}(x) - (M-x)\Pi_{x}^{+}(x)}$$

As

$$\frac{M - \Psi^+(x)}{M - x} = \frac{\int_x^M \Pi^+(y) dy}{(M - x)\Pi^+(x)} - 1,$$

we must have $\Psi_x^+(M) = \frac{\alpha}{\alpha+1} < 1$.

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